Package 'sparr'

July 7, 2024

```
Type Package
Title Spatial and Spatiotemporal Relative Risk
Version 2.3-15
Date 2024-07-06
Description Provides functions to estimate kernel-smoothed spatial and spatio-temporal densi-
     ties and relative risk functions, and perform subsequent inference. Methodological de-
     tails can be found in the accompanying tutorial: Davies et al. (2018) < DOI:10.1002/sim.7577>.
Depends R (>= 2.10.1), spatstat (>= 2.3-0)
Imports spatstat.utils, spatstat.geom, spatstat.explore,
     spatstat.random, spatstat.univar, doParallel, parallel,
     foreach, misc3d
Suggests fftwtools (>= 0.9.8)
License GPL (>= 2)
LazyLoad yes
NeedsCompilation no
RoxygenNote 7.3.2
URL https://tilmandavies.github.io/sparr/,
     https://github.com/tilmandavies/sparr/
BugReports https://github.com/tilmandavies/sparr/issues/
Encoding UTF-8
Author Tilman M. Davies [aut, cre],
     Jonathan C. Marshall [aut]
Maintainer Tilman M. Davies <tilman.davies@otago.ac.nz>
Repository CRAN
Date/Publication 2024-07-07 10:40:02 UTC
```

Contents

		he sparr Package: Sr	
Index			69
	tolerance		 66
	tol.contour		 65
	•		
	•		
	1		
			 46
			 45
			 44
	1		 41
	•		 40
	•		
	•		
			19
	1 1		 16
	•		 13
	•		 9
			 8
	1 1		 2

Description

Provides functions to estimate fixed and adaptive kernel-smoothed spatial relative risk surfaces via the density-ratio method and perform subsequent inference. Fixed-bandwidth spatiotemporal density and relative risk estimation is also supported.

Details

Package: sparr
Date: 2024-07-06
Version: 2.3-15
License: GPL (>= 2)

Kernel smoothing, and the flexibility afforded by this methodology, provides an attractive approach to estimating complex probability density functions.

The *spatial relative risk function*, constructed as a ratio of estimated case to control densities (Bithell, 1990; 1991; Kelsall and Diggle, 1995a,b), describes the variation in the 'risk' of the disease, given the underlying at-risk population. This is a technique that has been applied successfully for mainly exploratory purposes in a number of different analyses (see for example Sabel et al., 2000; Prince et al., 2001; Wheeler, 2007, Elson et al., 2021). It has also grown in popularity in very different fields that pose similarly styled research questions, such as ecology (e.g. Campos and Fedigan, 2014); physiology (Davies et al., 2013); and archaeology (e.g. Bevan, 2012; Smith et al. 2015).

This package provides functions for spatial (i.e. bivariate/planar/2D) kernel density estimation (KDE), implementing both fixed and 'variable' or 'adaptive' (Abramson, 1982) smoothing parameter options. A selection of bandwidth calculators for bivariate KDE and the relative risk function are provided, including one based on the maximal smoothing principle (Terrell, 1990), and others involving a leave-one-out cross-validation (see below). In addition, the ability to construct both Monte-Carlo and asymptotic *p*-value surfaces ('tolerance' contours of which signal statistically significant sub-regions of extremity in a risk surface - Hazelton and Davies, 2009; Davies and Hazelton, 2010) as well as some visualisation tools are provided.

Spatiotemporal estimation is also supported, largely following developments in Fernando and Hazelton (2014). This includes their fixed-bandwith kernel estimator of spatiotemporal densities, relative risk, and asymptotic tolerance contours.

Key content of sparr can be broken up as follows:

DATASETS/DATA GENERATION

pbc a case/control planar point pattern (ppp.object) concerning liver disease in northern England. fmd an anonymised (jittered) case/control spatiotemporal point pattern of the 2001 outbreak of veterinary foot-and-mouth disease in Cumbria (courtesy of the Animal and Plant Health Agency, UK).

burk a spatiotemporal point pattern of Burkitt's lymphoma in Uganda; artificially simulated control data are also provided for experimentation.

Also available are a number of relevant additional spatial datasets built-in to the spatstat package (Baddeley and Turner, 2005; Baddeley et al., 2015) through spatstat.data, such as chorley, which concerns the distribution of laryngeal cancer in an area of Lancashire, UK.

rimpoly a wrapper function of rpoint to allow generated spatial point patterns based on a pixel image to be returned with a polygonal owin.

SPATIAL

Bandwidth calculators

OS estimation of an isotropic smoothing parameter for fixed-bandwidth bivariate KDE, based on the oversmoothing principle introduced by Terrell (1990).

NS estimation of an isotropic smoothing parameter for fixed-bandwidth bivariate KDE, based on the asymptotically optimal value for a normal density (bivariate normal scale rule - see e.g. Wand and Jones, 1995).

LSCV. density a least-squares cross-validated (LSCV) estimate of an isotropic fixed bandwidth for bivariate, edge-corrected KDE (see e.g. Bowman and Azzalini, 1997).

LIK. density a likelihood cross-validated (LIK) estimate of an isotropic fixed bandwidth for bivariate, edge-corrected KDE (see e.g. Silverman, 1986).

SLIK. adapt an experimental likelihood cross-validation function for simultaneous global/pilot bandwidth selection for adaptive density estimates.

BOOT. density a bootstrap approach to optimisation of an isotropic fixed bandwidth for bivariate, edge-corrected KDE (see e.g. Taylor, 1989).

LSCV.risk Estimation of a jointly optimal, common isotropic case-control fixed bandwidth for the kernel-smoothed risk function based on the mean integrated squared error (MISE), a weighted MISE, or the asymptotic MISE (see respectively Kelsall and Diggle, 1995a; Hazelton, 2008; Davies, 2013).

Density and relative risk estimation

bivariate.density kernel density estimate of bivariate data; fixed or adaptive smoothing.

multiscale.density multi-scale adaptive kernel density estimates for multiple global bandwidths as per Davies and Baddeley (2018).

multiscale.slice a single adaptive kernel estimate based on taking a slice from a multi-scale estimate.

risk estimation of a (log) spatial relative risk function, either from data or pre-existing bivariate density estimates; fixed (Kelsall and Diggle, 1995a); fixed with shrinkage (Hazelton, 2023); or both asymmetric (Davies and Hazelton, 2010) and symmetric (Davies et al., 2016) adaptive estimates are possible.

tolerance calculation of asymptotic or Monte-Carlo *p*-value surfaces.

tol.classify classification of individual points based on a *p*-value surface from which tolerance contours are produced.

Visualisation

S3 methods of the plot function; see plot.bivden for visualising a single bivariate density estimate from bivariate.density, plot.rrs for visualisation of a spatial relative risk function from risk, or plot.msden for viewing animations of multi-scale density estimates from multiscale.density.

tol.contour provides more flexibility for plotting and superimposing tolerance contours upon an existing plot of spatial relative risk (i.e. given output from tolerance).

tol.classplot is a utility to view the classification scheme with respect to unique tolerance contours/significance regions as identified by tol.classify.

Printing and summarising

S3 methods (print.bivden, print.rrs, print.msden, summary.bivden, summary.rrs, and summary.msden) are available for the bivariate density, spatial relative risk, and multi-scale adaptive density objects.

SPATIOTEMPORAL

Bandwidth calculators

OS. spattemp estimation of an isotropic smoothing parameter for the spatial margin and another for the temporal margin for spatiotemporal densities, based on the 2D and 1D versions, respectively, of the oversmoothing principle introduced by Terrell (1990).

NS. spattemp as above, based on the 2D and 1D versions of the normal scale rule (Silverman, 1986).

LSCV. spattemp least-squares cross-validated (LSCV) estimates of scalar spatial and temporal bandwidths for edge-corrected spatiotemporal KDE.

LIK. spattemp as above, based on likelihood cross-validation.

BOOT. spattemp bootstrap bandwidth selection for the spatial and temporal margins; for spatiotemporal, edge-corrected KDE (Taylor, 1989).

Density and relative risk estimation

spattemp.density fixed-bandwidth kernel density estimate of spatiotemporal data.

spattemp.risk fixed-bandwidth kernel density estimate of spatiotemporal relative risk, either with a time-static or time-varying control density (Fernando and Hazelton, 2014).

spattemp.slice extraction function of the spatial density/relative risk at prespecified time(s).

Visualisation

S3 methods of the plot function; see plot.stden for various options (including animation) for visualisation of a spatiotemporal density, and plot.rrst for viewing spatiotemporal relative risk surfaces (including animation and tolerance contour superimposition).

Printing and summarising objects

S3 methods (print.stden, print.rrst, summary.stden, and summary.rrst) are available for the spatiotemporal density and spatiotemporal relative risk objects respectively.

Dependencies

The sparr package depends upon spatstat. In particular, the user should familiarise themselves with ppp objects and im objects, which are used throughout. For spatiotemporal density estimation, sparr is assisted by importing from the misc3d package, and for the experimental capabilities involving parallel processing, sparr also currently imports doParallel, parallel, and foreach.

Citation

To cite use of current versions of sparr in publications or research projects please use:

Davies, T.M., Marshall, J.C. and Hazelton, M.L. (2018) Tutorial on kernel estimation of continuous spatial and spatiotemporal relative risk, *Statistics in Medicine*, **37**(7), 1191-1221. <DOI:10.1002/sim.7577>

Old versions of sparr (<= 2.1-09) can be referenced by Davies et al. (2011) (see reference list).

Author(s)

T.M. Davies

Dept. of Mathematics & Statistics, University of Otago, Dunedin, New Zealand.

J.C. Marshall

Institute of Fundamantal Sciences, Massey University, Palmerston North, New Zealand.

Maintainer: T.M.D. <tilman.davies@otago.ac.nz>

References

Abramson, I. (1982), On bandwidth variation in kernel estimates — a square root law, *Annals of Statistics*, **10**(4), 1217-1223.

Baddeley, A. and Turner, R. (2005), spatstat: an R package for analyzing spatial point patterns, *Journal of Statistical Software*, **12**(6), 1-42.

Baddeley, A., Rubak, E. and Turner, R. (2015) *Spatial Point Patterns: Methodology and Applications with R*, Chapman and Hall/CRC Press, UK.

Bevan A. (2012), Spatial methods for analysing large-scale artefact inventories. *Antiquity*, **86**, 492-506.

Bithell, J.F. (1990), An application of density estimation to geographical epidemiology, *Statistics in Medicine*, **9**, 691-701.

Bithell, J.F. (1991), Estimation of relative risk function, Statistics in Medicine, 10, 1745-1751.

Bowman, A.W. and Azzalini, A. (1997), *Applied Smoothing Techniques for Data Analysis: The Kernel Approach with S-Plus Illustrations*. Oxford University Press Inc., New York. ISBN 0-19-852396-3.

Campos, F.A. and Fedigan, L.M. (2014) Spatial ecology of perceived predation risk and vigilance behavior in white-faced capuchins, *Behavioral Ecology*, **25**, 477-486.

Davies, T.M. (2013), Jointly optimal bandwidth selection for the planar kernel-smoothed density-ratio, *Spatial and Spatio-temporal Epidemiology*, **5**, 51-65.

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

Davies, T.M., Cornwall, J. and Sheard, P.W. (2013) Modelling dichotomously marked muscle fibre configurations, *Statistics in Medicine*, **32**, 4240-4258.

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Davies, T.M., Hazelton, M.L. and Marshall, J.C. (2011), sparr: Analyzing spatial relative risk using fixed and adaptive kernel density estimation in R, *Journal of Statistical Software* **39**(1), 1-14.

Davies, T.M., Jones, K. and Hazelton, M.L. (2016), Symmetric adaptive smoothing regimens for estimation of the spatial relative risk function, *Computational Statistics & Data Analysis*, **101**, 12-28.

Elson, R., Davies, T. M., Lake, I. R., Vivancos, R., Blomquist, P. B., Charlett, A. and Dabrera, G. (2021), The spatio-temporal distribution of COVID-19 infection in England between January and June 2020, *Epidemiology and Infection*, **149**, e73.

Fernando, W.T.P.S. and Hazelton, M.L. (2014), Generalizing the spatial relative risk function, *Spatial and Spatio-temporal Epidemiology*, **8**, 1-10.

Hazelton, M.L. (2008), Letter to the editor: Kernel estimation of risk surfaces without the need for edge correction, *Statistics in Medicine*, **27**, 2269-2272.

Hazelton, M.L. (2023), Shrinkage estimators of the spatial relative risk function, *Submitted for publication*.

Hazelton, M.L. and Davies, T.M. (2009), Inference based on kernel estimates of the relative risk function in geographical epidemiology, *Biometrical Journal*, **51**(1), 98-109.

Kelsall, J.E. and Diggle, P.J. (1995a), Kernel estimation of relative risk, Bernoulli, 1, 3-16.

Kelsall, J.E. and Diggle, P.J. (1995b), Non-parametric estimation of spatial variation in relative risk, *Statistics in Medicine*, **14**, 2335-2342.

Prince, M. I., Chetwynd, A., Diggle, P. J., Jarner, M., Metcalf, J. V. and James, O. F. W. (2001), The geographical distribution of primary biliary cirrhosis in a well-defined cohort, *Hepatology* **34**, 1083-1088.

Sabel, C. E., Gatrell, A. C., Loytonen, M., Maasilta, P. and Jokelainen, M. (2000), Modelling exposure opportunitites: estimating relative risk for motor disease in Finland, *Social Science & Medicine* **50**, 1121-1137.

Smith, B.A., Davies, T.M. and Higham, C.F.W. (2015) Spatial and social variables in the Bronze Age phase 4 cemetery of Ban Non Wat, Northeast Thailand, *Journal of Archaeological Science: Reports*, **4**, 362-370.

Taylor, C.C. (1989) Bootstrap choice of the smoothing parameter in kernel density estimation, *Biometrika*, **76**, 705-712.

Terrell, G.R. (1990), The maximal smoothing principle in density estimation, *Journal of the American Statistical Association*, **85**, 470-477.

Venables, W. N. and Ripley, B. D. (2002). *Modern Applied Statistics with S*, Fourth Edition, Springer, New York.

Wand, M.P. and Jones, C.M., 1995. Kernel Smoothing, Chapman & Hall, London.

Wheeler, D. C. (2007), A comparison of spatial clustering and cluster detection techniques for childhood leukemia incidence in Ohio, 1996-2003, *International Journal of Health Geographics*, **6**(13).

See Also

Useful links:

- https://tilmandavies.github.io/sparr/
- https://github.com/tilmandavies/sparr/
- Report bugs at https://github.com/tilmandavies/sparr/issues/

8 available.h0

available.h0

Available global bandwidth range

Description

Gets universally available global bandwidths as represented by several multi-scale density estimate objects

Usage

```
available.h0(...)
```

Arguments

... Any number of objects of class msden; possibly named.

Details

This simple function merely accesses and returns the maximum lower limit and minimum upper limit of all hørange components of the msden objects passed through Natural numeric error arising from any changes to the bandwidth-axis discretisation resolution in the creation of the msden objects (i.e. through the 'dimz' argument) means individual global bandwidth ranges can differ slightly between affected multi-scale estimates, even if they are all applied to the same data set. Can additionally be useful when, for example, creating asymmetric relative risk surfaces based on slices of multi-scale densities with respect to the case and control data sets, because the bandwidth factors differ.

Throws an error if one or more of the horange components is incompatible (i.e. all horange components must overlap).

Value

A numeric vector of length 2 providing the range of available global bandwidths compatible with all supplied multi-scale density estimates.

Author(s)

T.M. Davies

See Also

```
multiscale.density, multiscale.slice
```

Examples

```
# See ?multiscale.slice
```

bivariate.density Bivariate kernel density/intensity estimation

Description

Provides an isotropic adaptive or fixed bandwidth kernel density/intensity estimate of bivariate/planar/2D data.

Usage

```
bivariate.density(
  pp,
  h0,
  hp = NULL,
  adapt = FALSE,
  resolution = 128,
  gamma.scale = "geometric",
  edge = c("uniform", "diggle", "none"),
 weights = NULL,
  intensity = FALSE,
  trim = 5,
  xy = NULL,
  pilot.density = NULL,
  leaveoneout = FALSE,
  parallelise = NULL,
 davies.baddeley = NULL,
  verbose = TRUE
)
```

Arguments

рр	An object of class ppp giving the observed 2D data set to be smoothed.
h0	Global bandwidth for adaptive smoothing or fixed bandwidth for constant smoothing. A numeric value > 0 .
hp	Pilot bandwidth (scalar, numeric > 0) to be used for fixed bandwidth estimation of a pilot density in the case of adaptive smoothing. If NULL (default), it will take on the value of h0. Ignored when adapt = FALSE or if pilot.density is supplied as a pre-defined pixel image.
adapt	Logical value indicating whether to perform adaptive kernel estimation. See 'Details'.
resolution	Numeric value > 0 . Resolution of evaluation grid; the density/intensity will be returned on a [resolution \times resolution] grid.
gamma.scale	Scalar, numeric value > 0; controls rescaling of the variable bandwidths. Defaults to the geometric mean of the bandwidth factors given the pilot density (as per Silverman, 1986). See 'Details'.

character string giving the type of edge correction to employ. "uniform" (default) corrects based on evaluation grid coordinate and "diggle" reweights each observation-specific kernel. Setting edge = "none" requests no edge correction. Further details can be found in the documentation for density.ppp.

Optional numeric vector of nonnegative weights corresponding to each observation in pp. Must have length equal to npoints(pp).

Logical value indicating whether to return an intensity estimate (integrates to the sample size over the study region), or a density estimate (default, integrates

to 1).

trim Numeric value > 0; controls bandwidth truncation for adaptive estimation. See

'Details'.

xy Optional alternative specification of the evaluation grid; matches the argument

of the same tag in as.mask. If supplied, resolution is ignored.

pilot.density An optional pixel image (class im) giving the pilot density to be used for cal-

culation of the variable bandwidths in adaptive estimation, **or** a ppp.object giving the data upon which to base a fixed-bandwidth pilot estimate using hp. If used, the pixel image *must* be defined over the same domain as the data given resolution or the supplied pre-set xy evaluation grid; **or** the planar point pattern data must be defined with respect to the same polygonal study region as in

pp.

leaveoneout Logical value indicating whether to compute and return the value of the den-

sity/intensity at each data point for an adaptive estimate. See 'Details'.

parallelise Numeric argument to invoke parallel processing, giving the number of CPU

cores to use when leaveoneout = TRUE. Experimental. Test your system first using parallel::detectCores() to identify the number of cores available to

you.

davies.baddeley

weights

intensity

An optional numeric vector of length 3 to control bandwidth partitioning for approximate adaptive estimation, giving the quantile step values for the variable bandwidths for density/intensity and edge correction surfaces and the resolution of the edge correction surface. May also be provided as a single numeric value.

See 'Details'.

verbose Logical value indicating whether to print a function progress bar to the console

when adapt = TRUE.

Details

Given a data set x_1, \ldots, x_n in 2D, the isotropic kernel estimate of its probability density function, $\hat{f}(x)$, is given by

$$\hat{f}(y) = n^{-1} \sum_{i=1}^{n} h(x_i)^{-2} K((y - x_i)/h(x_i))$$

where h(x) is the bandwidth function, and K(.) is the bivariate standard normal smoothing kernel. Edge-correction factors (not shown above) are also implemented.

Fixed The classic fixed bandwidth kernel estimator is used when adapt = FALSE. This amounts to setting h(u) = h0 for all u. Further details can be found in the documentation for density.ppp.

Adaptive Setting adapt = TRUE requests computation of Abramson's (1982) variable-bandwidth estimator. Under this framework, we have $h(u) = h0*min[\tilde{f}(u)^{-1/2}, G*trim]/\gamma$, where $\tilde{f}(u)$ is a fixed-bandwidth kernel density estimate computed using the pilot bandwidth hp.

- Global smoothing of the variable bandwidths is controlled with the global bandwidth h0.
- In the above statement, G is the geometric mean of the "bandwidth factors" $\tilde{f}(x_i)^{-1/2}$; $i=1,\ldots,n$. By default, the variable bandwidths are rescaled by $\gamma=G$, which is set with gamma.scale = "geometric". This allows h0 to be considered on the same scale as the smoothing parameter in a fixed-bandwidth estimate i.e. on the scale of the recorded data. You can use any other rescaling of h0 by setting gamma.scale to be any scalar positive numeric value; though note this only affects γ see the next bullet. When using a scale-invariant h0, set gamma.scale = 1.
- The variable bandwidths must be trimmed to prevent excessive values (Hall and Marron, 1988). This is achieved through trim, as can be seen in the equation for h(u) above. The trimming of the variable bandwidths is universally enforced by the geometric mean of the bandwidth factors G independent of the choice of γ . By default, the function truncates bandwidth factors at five times their geometric mean. For stricter trimming, reduce trim, for no trimming, set trim = Inf.
- For even moderately sized data sets and evaluation grid resolution, adaptive kernel estimation can be rather computationally expensive. The argument davies baddeley is used to approximate an adaptive kernel estimate by a sum of fixed bandwidth estimates operating on appropriate subsets of pp. These subsets are defined by "bandwidth bins", which themselves are delineated by a quantile step value $0 < \delta < 1$. E.g. setting $\delta = 0.05$ will create 20 bandwidth bins based on the 0.05th quantiles of the Abramson variable bandwidths. Adaptive edge-correction also utilises the partitioning, with pixelwise bandwidth bins defined using the value $0 < \beta < 1$, and the option to decrease the resolution of the edge-correction surface for computation to a $[L \times L]$ grid, where $0 < L \le$ resolution. If davies baddeley is supplied as a vector of length 3, then the values [1], [2], and [3] correspond to the parameters δ , β , and $L_M = L_N$ in Davies and Baddeley (2018). If the argument is simply a single numeric value, it is used for both δ and β , with $L_M = L_N$ =resolution (i.e. no edge-correction surface coarsening).
- Computation of leave-one-out values (when leaveoneout = TRUE) is done by brute force, and is therefore very computationally expensive for adaptive smoothing. This is because the leave-one-out mechanism is applied to both the pilot estimation and the final estimation stages. Experimental code to do this via parallel processing using the foreach routine is implemented. Fixed-bandwidth leave-one-out can be performed directly in density.ppp.

Value

If leaveoneout = FALSE, an object of class "bivden". This is effectively a list with the following components:

z The resulting density/intensity estimate, a pixel image object of class im.

h0 A copy of the value of h0 used.

hp A copy of the value of hp used.

h A numeric vector of length equal to the number of data points, giving the bandwidth used for the corresponding observation in pp.

him	A pixel image (class im), giving the 'hypothetical' Abramson bandwidth at each pixel coordinate conditional upon the observed data. NULL for fixed-bandwidth estimates.
q	Edge-correction weights; a pixel image if edge = "uniform", a numeric vector if edge = "diggle", and NULL if edge = "none".
gamma	The value of γ used in scaling the bandwidths. NA if a fixed bandwidth estimate is computed.
geometric	The geometric mean G of the untrimmed bandwidth factors $\tilde{f}(x_i)^{-1/2}$. NA if a fixed bandwidth estimate is computed.
рр	A copy of the ppp.object initially passed to the pp argument, containing the data that were smoothed.

Else, if leaveoneout = TRUE, simply a numeric vector of length equal to the number of data points, giving the leave-one-out value of the function at the corresponding coordinate.

Author(s)

T.M. Davies and J.C. Marshall

References

Abramson, I. (1982). On bandwidth variation in kernel estimates — a square root law, *Annals of Statistics*, **10**(4), 1217-1223.

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Davies, T.M., Jones, K. and Hazelton, M.L. (2016), Symmetric adaptive smoothing regimens for estimation of the spatial relative risk function, *Computational Statistics & Data Analysis*, **101**, 12-28.

Diggle, P.J. (1985), A kernel method for smoothing point process data, *Journal of the Royal Statistical Society, Series C*, **34**(2), 138-147.

Hall P. and Marron J.S. (1988) Variable window width kernel density estimates of probability densities. *Probability Theory and Related Fields*, **80**, 37-49.

Marshall, J.C. and Hazelton, M.L. (2010) Boundary kernels for adaptive density estimators on regions with irregular boundaries, *Journal of Multivariate Analysis*, **101**, 949-963.

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

Wand, M.P. and Jones, C.M., 1995. Kernel Smoothing, Chapman & Hall, London.

Examples

```
data(chorley) # Chorley-Ribble data from package 'spatstat'
# Fixed bandwidth kernel density; uniform edge correction
chden1 <- bivariate.density(chorley,h0=1.5)</pre>
```

BOOT.density 13

```
# Fixed bandwidth kernel density; diggle edge correction; coarser resolution
chden2 <- bivariate.density(chorley,h0=1.5,edge="diggle",resolution=64)

# Adaptive smoothing; uniform edge correction
chden3 <- bivariate.density(chorley,h0=1.5,hp=1,adapt=TRUE)

# Adaptive smoothing; uniform edge correction; partitioning approximation
chden4 <- bivariate.density(chorley,h0=1.5,hp=1,adapt=TRUE,davies.baddeley=0.025)

oldpar <- par(mfrow=c(2,2))
plot(chden1);plot(chden2);plot(chden3);plot(chden4)
par(oldpar)</pre>
```

BOOT.density

Bootstrap bandwidth for a spatial kernel density estimate

Description

Isotropic fixed or global (for adaptive) bandwidth selection for a standalone 2D density based on bootstrap estimation of the MISE.

Usage

```
BOOT.density(pp, hlim = NULL, eta = NULL, type = c("fixed", "adaptive"), hp = NULL, edge = c("uniform", "none"), ref.density = NULL, resolution = 64, rmdiag = TRUE, sim.adapt = list(N = 50, B = 100, dimz = 64, objective = FALSE), parallelise = NA, verbose = TRUE, ...)
```

Arguments

pp	An object of class ppp giving the observed 2D data to be smoothed.
hlim	An optional vector of length 2 giving the limits of the optimisation routine with respect to the bandwidth. If NULL, the function attempts to choose this automatically.
eta	Fixed scalar bandwidth to use for the reference density estimate; if NULL it is calculated as the oversmoothing bandwidth of pp using OS. Ignored if ref.density is supplied. See 'Details'.
type	A character string indicating selection type. Either "fixed" (default) for selection of a constant bandwidth for the fixed-bandwidth estimator based on theory extended from results in Taylor (1989); or "adaptive" for selection of the global bandwidth for an adaptive kernel density. See 'Details'.
hp	Pilot bandwidth used for adaptive estimates in the bootstrap; see the argument of the same tag in bivariate.density. Ignored when type = "fixed" or when ref.density is supplied.

14 BOOT.density

edge	Character string dictating edge correction for the bootstrapped estimates. "uniform" (default) corrects based on evaluation grid coordinate. Setting edge="none" requests no edge correction.
ref.density	Optional. An object of class bivden giving the reference density from which data will be generated. Based on theory, this must be a fixed-bandwidth estimate if type = "fixed"; see 'Details'. Must be edge-corrected if edge = "uniform".
resolution	Spatial grid size; the optimisation will be based on a [resolution \times resolution] density estimate.
rmdiag	Logical control value for removal of mirrored evaluation points as suggested by Taylor (1989) in the theoretical expression of the fixed-bandwidth MISE estimate. See 'Details'. Ignored when type = "adaptive"
sim.adapt	List of control values for bootstrap simulation in the adaptive case; see 'Details'. Ignored when type = "fixed".
parallelise	Optional numeric argument to reduce computation time by invoking parallel processing, by giving the number of CPU cores to use in either evaluation (fixed) or in the actual bootstrap replicate generation (adaptive). Experimental. Test your system first using parallel::detectCores() to identify the number of cores available to you.
verbose	Logical value indicating whether to print function progress during execution.
	Optional arguments controlling scaling to be passed to multiscale.density for the adaptive bootstrap; ignored when type = "fixed".

Details

For a 2D kernel density estimate \hat{f} defined on $W \in \mathbb{R}^2$, the mean integrated squared error (MISE) is given by $E[\int_W (\hat{f}(x) - f(x))^2 dx]$, where f is the corresponding true density. Given an observed data set X (argument pp) of n observations, this function finds the bandwidth h that minimises

$$E^* [\int_W (\hat{f}^*(x) - \hat{f}(x))^2 dx],$$

where $\hat{f}(x)$ is a density estimate of X constructed with 'reference' bandwidth η (argument eta or ref.density), and $\hat{f}^*(x)$ is a density estimate using bandwidth h of n observations X^* generated from $\hat{f}(x)$. The notation E^* denotes expectation with respect to the distribution of the X^* .

Fixed When type = "fixed", the function assumes you want to select a constant bandwidth for use with the fixed-bandwith density estimator. This implementation is based on extending the remarkable results of Taylor (1989) (see also Sain et al., 1994), who demonstrates that when the Gaussian kernel is being used, we can find the optimal h with respect to the aforementioned bootstrap-estimated MISE without any actual resampling. This implementation extends these results to the bivariate setting, and allows for edge-correction of both the reference and bootstrap densities.

• Taylor (1989) does not distinguish between the reference bandwidth η and the target of optimisation, h, thus allowing the reference bandwidth to vary alongside the target in the optimisation. This is not optimal, and this function always assumes a static reference bandwidth. Hall et al. (1992) indicate that a generous amount of smoothing is to be preferred in the reference density (hence the default eta set using 0S).

BOOT.density 15

• If ref.density is supplied, it **must** be a fixed-bandwidth density estimate as an object of class bivden for validity of the theory. Edge-correction must be present if edge = "uniform"; and it must be evaluated on the same spatial domain as dictated by Window(pp) and resolution. If unsupplied, the function internally computes an appropriate fixed-bandwidth density estimate using eta as the reference bandwidth.

• Finally, Taylor (1989) argues it is preferable to avoid summation at identical evaluation grid points in the expression for the optimal bandwidth, which is performed when rmdiag = TRUE. Setting rmdiag = FALSE disables this correction.

Adaptive When type = "adaptive", the function assumes you want to select a global bandwidth (argument h0 in bivariate.density) for use in 2D adaptive kernel density estimation.

- An expression similar to Taylor (1989) is not possible for the adaptive estimator. Thus, in the adaptive setting, the optimal bootstrap bandwidth is calculated by brute force as was performed in Davies and Baddeley (2018) by taking advantage of the multiscale estimation theory implemented in multiscale.density. The value that minimises an interpolating cubic spline of the estimated MISE on bandwidth is identified as the optimal global bandwidth.
- The user can pass either a fixed or adaptive bivden object to ref. density. If this is the case, hp is ignored and the pilot bandwidth for each iteration of the bootstrap in estimation of the $\hat{f}^*(x)$ uses ref.density\$hp (if ref.density is adaptive) or ref.density\$h0 (if ref.density is fixed). When ref.density is unsupplied, the function uses a fixed-bandwidth kernel estimate with bandwidth eta as the reference density, and if additionally hp is unsupplied, the same value eta is used for the constant pilot bandwidth.
- Control over the bootstrap is achieved with four optional named arguments passed as a list to sim.adapt. N controls the number of bootstrap iterates per bandwidth; B controls the resolution of the sequence of bandwidths trialled (i.e. between hlim[1] and hlim[2]); dimz specifies the resolution of the bandwidth axis in the trivariate convolution evaluated by multiscale.density; and objective specifies whether to return the set of estimated MISEs for all bandwidths (nice to plot), or merely the optimal bandwidth (see 'Value').
- The ... are intended for any relevant optional arguments to be passed to the internal call to multiscale.density, such as gamma.scale or trim.

Value

The optimal fixed or global (for adaptive) scalar bandwidth. If simargs\$objective = TRUE for the adaptive bootstrap, the return object is instead a [simargs\$B x2] matrix, with the first column giving the trialled bandwidth and the second giving the corresponding value of the estimated bootstrap MISE.

Warning

Even with the implemented computational tricks, bootstrapping for bandwidth selection for spatial data is still computationally demanding, especially for adaptive kernel estimates. The user can reduce this time by keeping the evaluation grid at modest resolutions, and experimenting with parallelising the internal loops via parallelise. The 'Examples' section offers some rough indications of evaluation times on this author's local machine.

Author(s)

T.M. Davies

16 BOOT.spattemp

References

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

Hall, P., Marron, J.S. and Park, B.U. (1992) Smoothed cross-validation, *Probability Theory and Related Fields*, **92**, 1-20.

Sain, S.R., Baggerly, K.A. and Scott, D.W. (1994) Cross-validation of multivariate densities, *Journal of the American Statistical Association*, **89**, 807-817.

Taylor, C.C. (1989) Bootstrap choice of the smoothing parameter in kernel density estimation, *Biometrika*, **76**, 705-712.

See Also

bivariate.density, OS, multiscale.density

Examples

```
data(pbc)
## Fixed bandwidth selection ##
BOOT.density(pbc) # ~20 secs
BOOT.density(pbc,eta=OS(pbc)/2) # halve default reference bandwidth
BOOT.density(pbc,eta=OS(pbc)*2) # double default reference bandwidth
# supplying pre-defined reference density as fixed-bandwidth 'bivden' object
pbcfix <- bivariate.density(pbc,h0=2.5,resolution=64)</pre>
system.time(hfix <- BOOT.density(pbc,ref.density=pbcfix))</pre>
hfix
## Global (for adaptive) bandwidth selection ##
# ~200 secs next line; use 'parallelise' for speedup
system.time(hada <- BOOT.density(pbc,type="adaptive")) # minimal usage for adaptive bootstrap
hada
# ~80 secs next line. Set custom h limits; increase reference bandwidth;
    set custom pilot bandwidth; return objective function; use 'parallelise' for speedup
system.time(hada <- BOOT.density(pbc,hlim=c(0.9,8),eta=3.5,type="adaptive",
                                 hp=OS(pbc)/2, sim. adapt=list(objective=TRUE)))
hada[which.min(hada[,2]),1]
plot(hada);abline(v=hada[which.min(hada[,2]),1],col=2)
```

BOOT.spattemp 17

Description

Bandwidth selection for standalone spatiotemporal density/intensity based on bootstrap estimation of the MISE, providing an isotropic scalar spatial bandwidth and a scalar temporal bandwidth.

Usage

```
BOOT.spattemp(pp, tt = NULL, tlim = NULL, eta = NULL, nu = NULL, sedge = c("uniform", "none"), tedge = sedge, ref.density = NULL, sres = 64, tres = sres, start = NULL, verbose = TRUE)
```

Arguments

pp	An object of class ppp giving the spatial coordinates of the observations to be smoothed. Possibly marked with the time of each event; see argument tt.
tt	A numeric vector of equal length to the number of points in pp, giving the time corresponding to each spatial observation. If unsupplied, the function attempts to use the values in the marks attribute of the ppp.object in pp.
tlim	A numeric vector of length 2 giving the limits of the temporal domain over which to smooth. If supplied, all times in tt must fall within this interval (equality with limits allowed). If unsupplied, the function simply uses the range of the observed temporal values.
eta	Fixed scalar bandwidth to use for the spatial margin of the reference density estimate; if NULL it is calculated as the oversmoothing bandwidth of pp using OS. Ignored if ref.density is supplied. See 'Details'.
nu	Fixed scalar bandwidth to use for the temporal margin of the reference density estimate; if NULL it is calculated from tt using the univariate version of Terrell's (1990) oversmoothing principle. Ignored if ref. density is supplied. See 'Details'.
sedge	Character string dictating spatial edge correction. "uniform" (default) corrects based on evaluation grid coordinate. Setting sedge="none" requests no edge correction.
tedge	As sedge, for temporal edge correction.
ref.density	Optional. An object of class stden giving the reference density from which data is assumed to originate in the bootstrap. Must be spatially edge-corrected if sedge = "uniform".
sres	Numeric value > 0. Resolution of the [sres \times sres] evaluation grid in the spatial margin.
tres	Numeric value > 0. Resolution of the evaluation points in the temporal margin as defined by the tlim interval. If unsupplied, the density is evaluated at integer values between tlim[1] and tlim[2].
start	Optional positive numeric vector of length 2 giving starting values for the internal call to optim, in the order of (<spatial bandwidth="">, <temporal bandwidth="">).</temporal></spatial>
verbose	Logical value indicating whether to print a function progress bar to the console during evaluation.

18 BOOT.spattemp

Details

For a spatiotemporal kernel density estimate \hat{f} defined on $WxT \in R^3$, the mean integrated squared error (MISE) is given by $E[\int_W \int_T (\hat{f}(x,t) - f(x,t))^2 dt dx]$, where f is the corresponding true density. Given observed spatiotemporal locations X (arguments pp and tt) of n observations, this function finds the scalar spatial bandwidth h and scalar temporal bandwidth h that jointly minimise

$$E^* [\int_W \int_T (\hat{f}^*(x,t) - \hat{f}(x,t))^2 dt dx],$$

where $\hat{f}(x,t)$ is a density estimate of X constructed with 'reference' bandwidths η (spatial; argument eta) and ν (temporal; argument nu); $\hat{f}^*(x,t)$ is a density estimate using bandwidths h and λ of n observations X^* generated from $\hat{f}(x,t)$. The notation E^* denotes expectation with respect to the distribution of the X^* . The user may optionally supply ref.density as an object of class stden, which must be evaluated on the same spatial and temporal domains W and T as the data (arguments pp, tt, and tlim). In this case, the reference bandwidths are extracted from this object, and eta and nu are ignored.

This function is based on an extension of the theory of Taylor (1989) to the spatiotemporal domain and to cope with the inclusion of edge-correction factors. No resampling is necessary due to the theoretical properties of the Gaussian kernel.

Value

A numeric vector of length 2 giving the jointly optimised spatial and temporal bandwidths (named h and lambda respectively).

Warning

Bootstrapping for spatiotemporal bandwidth selection for spatiotemporal data is very computationally demanding. Keeping verbose = TRUE offers an indication of the computational burden by printing each pair of bandwidths at each iteration of the optimisation routine. The 'Examples' section also offers some rough indications of evaluation times on this author's local machine.

Author(s)

T. M. Davies

References

Taylor, C.C. (1989) Bootstrap choice of the smoothing parameter in kernel density estimation, *Biometrika*, **76**, 705-712.

See Also

LSCV.spattemp, spattemp.density

burk 19

Examples

```
data(burk) # Burkitt's Uganda lymphoma data
burkcas <- burk$cases

#~85 secs
hlam1 <- BOOT.spattemp(burkcas)

#~75 secs. Widen time limits, reduce ref. bw.
hlam2 <- BOOT.spattemp(burkcas,tlim=c(400,5800),eta=8,nu=450)

#~150 secs. Increase ref. bw., custom starting vals
hlam3 <- BOOT.spattemp(burkcas,eta=20,nu=800,start=c(7,400))

rbind(hlam1,hlam2,hlam3)</pre>
```

burk

Burkitt's lymphoma in Uganda

Description

Data of the spatiotemporal locations of Burkitt's lymphoma in the Western Nile district of Uganda from 1960 to 1975.

Format

burk is a named list with three members:

\$cases An object of class ppp giving the spatial locations (eastings/northings) of the 188 cases of Burkitt's lymphoma recorded in individuals of various ages (mostly children); the spatial study region as a polygonal owin; as well as the time (in days since 1/1/1960) of each observation stored as the marks of the points.

\$cases.age A numeric vector of length 188 giving the age of each individual in \$cases.

\$controls An object of class ppp giving 500 artificially simulated spatial-only observations to pose as a 'control' data set representing the at-risk population. The data were generated from a smooth kernel estimate of the spatial margin of the cases. The similarity between the case point distribution and the true at-risk population dispersion can be seen in e.g. Figure 2 of Middleton and Greenland (1954).

Source

The case data were extracted from the burkitt object of the splanes R package; see

Rowlingson B. and Diggle P.J. (2017), splancs: Spatial and Space-Time Point Pattern Analysis, R package version 2.01-40; https://CRAN.R-project.org/package=splancs.

20 fft2d

References

Bailey, T.C. and Gatrell, A.C. (1995), Interactive spatial data analysis, Longman; Harlow.

Middleton, J.F.M. and Greenland, D.J. (1954), Land and population in West Nile District, Uganda, *The Geographical Journal*, **120**, 446–455.

Examples

```
data(burk)
summary(burk$cases)

oldpar <- par(mfrow=c(1,3))
plot(burk$cases)
plot(burk$controls)
plot(density(marks(burk$cases)),xlim=range(marks(burk$cases)))
par(oldpar)</pre>
```

fft2d

2D fast-Fourier wrapper around 'fftwtools' or 'stats' package

Description

Utilises the Fastest Fourier Transform in the West (FFTW) via the 'fftwtools' package if available, else reverts to built-in functionality

Usage

```
fft2d(x, inverse = FALSE, fftw = fftw_available())
```

Arguments

A numeric matrix to be transformed.

inverse Whether it should compute the inverse transform (defaults to FALSE).

fftw Whether the fftwtools R package is available.

Details

This function is called wherever sparr seeks to perform a 2D fast-Fourier transform. Where available, computational expense is noticeably reduced by appealing to routines in the independent 'FFTW' toolbox. The user is encouraged to install the corresponding R package fftwtools from CRAN; this function will automatically detect and use the faster option, otherwise will defer to the built-in fft.

Value

The fast-Fourier (inverse) transform. A complex-valued matrix of the same size as x.

fmd 21

Author(s)

J.C. Marshall

Examples

```
# System check
sparr:::fftw_available()

system.time(fft(matrix(1:2000^2,2000)))
system.time(fft2d(matrix(1:2000^2,2000)))
```

fmd

Veterinary foot-and-mouth disease outbreak data

Description

Data of the spatial locations and time of farms infected by veterinary foot-and-mouth disease in the county of Cumbria, UK, over a course of nearly 250 days between February and August in 2001. There are 410 infected farms (the cases), and 1866 uninfected farms (the controls). The data have been jittered and randomly thinned by an unspecified amount to preserve anonymity.

Format

fmd is a named list with two members:

\$cases An object of class ppp giving the spatial locations of the 410 infected farms within a polygonal study region representing the county of Cumbria. The marks component of this object contain the integer day of infection (from beginning of study period).

\$controls An object of class ppp defined over the same spatial study region with the locations of the 1866 uninfected farms.

Acknowledgements

The Animal and Plant Health Agency (APHA), UK, provided permission to use this dataset.

References

Fernando, W.T.P.S. and Hazelton, M.L. (2014), Generalizing the spatial relative risk function, *Spatial and Spatio-temporal Epidemiology*, **8**, 1-10.

Keeling M, Woolhouse M, Shaw D, Matthews L, Chase-Topping M, Haydon D, et al. (2001), Dynamics of the 2001 UK foot and mouth epidemic: stochastic dispersal in a heterogeneous landscape, *Science*, **294**, 813-817.

Lawson A, Zhou H. (2005), Spatial statistical modeling of disease outbreaks with particular reference to the UK foot and mouth disease (FMD) epidemic of 2001, *Preventative Veterinary Medicine*, **71**, 141-156.

22 LIK.density

Examples

```
data(fmd)
summary(fmd$cases)
summary(fmd$controls)

oldpar <- par(mfrow=c(1,2))
plot(fmd$cases)
plot(fmd$controls)
par(oldpar)</pre>
```

LIK.density

Cross-validation bandwidths for spatial kernel density estimates

Description

Isotropic fixed or global (for adaptive) bandwidth selection for standalone 2D density/intensity based on either unbiased least squares cross-validation (LSCV) or likelihood (LIK) cross-validation.

Usage

```
LIK.density(
  pp,
  hlim = NULL,
  hseq = NULL,
  resolution = 64,
  edge = TRUE,
  auto.optim = TRUE,
  type = c("fixed", "adaptive"),
  seqres = 30,
  parallelise = NULL,
  zero.action = 0,
  verbose = TRUE,
)
LSCV.density(
  pp,
  hlim = NULL,
  hseq = NULL,
  resolution = 64,
  edge = TRUE,
  auto.optim = TRUE,
  type = c("fixed", "adaptive"),
  segres = 30,
  parallelise = NULL,
  zero.action = 0,
```

LIK.density 23

```
verbose = TRUE,
...
)
```

Arguments

pp An object of class ppp giving the observed 2D data to be smoothed.

hlim An optional vector of length 2 giving the limits of the optimisation routine with

respect to the bandwidth. If unspecified, the function attempts to choose this

automatically.

hseq An optional increasing sequence of bandwidth values at which to manually

evaluate the optimisation criterion. Used only in the case (!auto.optim &&

is.null(hlim)).

resolution Spatial grid size; the optimisation will be based on a [resolution × resolution]

density estimate.

edge Logical value indicating whether to edge-correct the density estimates used.

auto.optim Logical value indicating whether to automate the numerical optimisation using

optimise. If FALSE, the optimisation criterion is evaluated over hseq (if sup-

plied), or over a sequence of values controlled by hlim and seqres.

type A character string; "fixed" (default) performs classical leave-one-out cross-

validation for the fixed-bandwidth estimator. Alternatively, "adaptive" utilises multiscale adaptive kernel estimation (Davies & Baddeley, 2018) to run the cross-validation in an effort to find a suitable global bandwidth for the adaptive estimator. Note that data points are not 'left out' of the pilot density estimate when using this option (this capability is currently in development). See also the

entry for

seqres Optional resolution of an increasing sequence of bandwidth values. Only used

if (!auto.optim && is.null(hseq)).

parallelise Numeric argument to invoke parallel processing, giving the number of CPU

cores to use when !auto.optim and type = "fixed". Experimental. Test your system first using parallel::detectCores() to identify the number of cores

available to you.

zero.action A numeric integer, either -1, 0 (default), 1 or 2 controlling how the function

should behave in response to numerical errors at very small bandwidths, when such a bandwidth results in one or more zero or negative density values during

the leave-one-out computations. See 'Details'.

verbose Logical value indicating whether to provide function progress commentary.

.. Additional arguments controlling pilot density estimation and multi-scale bandwidth-

axis resolution when type = "adaptive". Relevant arguments are hp, pilot.density, gamma.scale, and trim from bivariate.density; and dimz from multiscale.density.

If hp is missing and required, the function makes a (possibly recursive) call to LSCV. density to set this using fixed-bandwidth LSCV. The remaining defaults are pilot.density = pp, gamma.scale = "geometric", trim = 5, and dimz =

resolution.

Details

This function implements the bivariate, edge-corrected versions of fixed-bandwidth least squares cross-validation and likelihood cross-validation as outlined in Sections 3.4.3 and 3.4.4 of Silverman (1986) in order to select an optimal fixed smoothing bandwidth. With type = "adaptive" it may also be used to select the global bandwidth for adaptive kernel density estimates, making use of multi-scale estimation (Davies and Baddeley, 2018) via multi-scale.density. Note that for computational reasons, the leave-one-out procedure is not performed on the pilot density in the adaptive setting; it is only performed on the final stage estimate. Current development efforts include extending this functionality, see SLIK.adapt. See also 'Warning' below.

Where LSCV.density is based on minimisation of an unbiased estimate of the mean integrated squared error (MISE) of the density, LIK.density is based on maximisation of the cross-validated leave-one-out average of the log-likelihood of the density estimate with respect to h.

In both functions, the argument zero.action can be used to control the level of severity in response to small bandwidths that result (due to numerical error) in at least one density value being zero or less. When zero.action = -1, the function strictly forbids bandwidths that would result in one or more *pixel* values of a kernel estimate of the original data (i.e. anything over the whole region) being zero or less—this is the most restrictive truncation. With zero.action = 0 (default), the function automatically forbids bandwidths that yield erroneous values at the leave-one-out data point locations only. With zero.action = 1, the minimum machine value (see .Machine\$double.xmin at the prompt) is used to replace these individual leave-one-out values. When zero.action = 2, the minimum value of the valid (greater than zero) leave-one-out values is used to replace any erroneous leave-one-out values.

Value

A single numeric value of the estimated bandwidth (if auto.optim = TRUE). Otherwise, a [seqres x 2] matrix giving the bandwidth sequence and corresponding CV function value.

Warning

Leave-one-out CV for bandwidth selection in kernel density estimation is notoriously unstable in practice and has a tendency to produce rather small bandwidths, particularly for spatial data. Satisfactory bandwidths are not guaranteed for every application; zero.action can curb adverse numeric effects for very small bandwidths during the optimisation procedures. This method can also be computationally expensive for large data sets and fine evaluation grid resolutions. The user may also need to experiment with adjusting hlim to find a suitable minimum.

Author(s)

T. M. Davies

References

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

Wand, M.P. and Jones, C.M., 1995. Kernel Smoothing, Chapman & Hall, London.

LIK.spattemp 25

See Also

SLIK.adapt and functions for bandwidth selection in package spatstat: bw.diggle; bw.ppl; bw.scott; bw.frac.

Examples

```
data(pbc)
pbccas <- split(pbc)$case</pre>
LIK.density(pbccas)
LSCV.density(pbccas)
#* FIXED
# custom limits
LIK.density(pbccas,hlim=c(0.01,4))
LSCV.density(pbccas,hlim=c(0.01,4))
# disable edge correction
LIK.density(pbccas,hlim=c(0.01,4),edge=FALSE)
LSCV.density(pbccas,hlim=c(0.01,4),edge=FALSE)
# obtain objective function
hcv <- LIK.density(pbccas,hlim=c(0.01,4),auto.optim=FALSE)</pre>
plot(hcv); abline(v=hcv[which.max(hcv[,2]),1],lty=2,col=2)
#* ADAPTIVE
LIK.density(pbccas,type="adaptive")
LSCV.density(pbccas,type="adaptive")
# change pilot bandwidth used
LIK.density(pbccas, type="adaptive", hp=2)
LSCV.density(pbccas,type="adaptive",hp=2)
```

LIK.spattemp

Cross-validation bandwidths for spatiotemporal kernel density estimates

Description

Bandwidth selection for standalone spatiotemporal density/intensity based on either unbiased least squares cross-validation (LSCV) or likelihood (LIK) cross-validation, providing an isotropic scalar spatial bandwidth and a scalar temporal bandwidth.

26 LIK.spattemp

Usage

```
LIK.spattemp(pp, tt = NULL, tlim = NULL, sedge = c("uniform", "none"),
  tedge = sedge, parallelise = NA, start = NULL, verbose = TRUE)

LSCV.spattemp(pp, tt = NULL, tlim = NULL, sedge = c("uniform", "none"),
  tedge = sedge, sres = 64, tres = sres, parallelise = NA,
  start = NULL, verbose = TRUE)
```

Arguments

рр	An object of class ppp giving the spatial coordinates of the observations to be smoothed. Possibly marked with the time of each event; see argument tt.
tt	A numeric vector of equal length to the number of points in pp, giving the time corresponding to each spatial observation. If unsupplied, the function attempts to use the values in the marks attribute of the ppp.object in pp.
tlim	A numeric vector of length 2 giving the limits of the temporal domain over which to smooth. If supplied, all times in tt must fall within this interval (equality with limits allowed). If unsupplied, the function simply uses the range of the observed temporal values.
sedge	Character string dictating spatial edge correction. "uniform" (default) corrects based on evaluation grid coordinate. Setting sedge="none" requests no edge correction.
tedge	As sedge, for temporal edge correction.
sres	Numeric value > 0 . Resolution of the [sres \times sres] evaluation grid in the spatial margin.
tres	Numeric value > 0. Resolution of the evaluation points in the temporal margin as defined by the tlim interval. If unsupplied, the density is evaluated at integer values between tlim[1] and tlim[2].
parallelise	Optional numeric argument to invoke parallel processing, by giving the number of CPU cores to use optimisation. This is only useful for larger data sets of many thousand observations. Experimental. Test your system first using parallel::detectCores() to identify the number of cores available to you.
start	Optional positive numeric vector of length 2 giving starting values for the internal call to optim, in the order of (<spatial bandwidth="">, <temporal bandwidth="">).</temporal></spatial>
verbose	Logical value indicating whether to print a function progress bar to the console during evaluation.

Value

A numeric vector of length 2 giving the jointly optimised spatial and temporal bandwidths (named h and lambda respectively).

Warning

Leave-one-out CV for bandwidth selection in kernel density estimation is notoriously unstable in practice and has a tendency to produce rather small bandwidths in the fixed bandwidth case. Satis-

factory bandwidths are not guaranteed for every application. This method can also be computationally expensive for large data sets and fine evaluation grid resolutions.

Author(s)

T. M. Davies

References

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

See Also

```
BOOT.spattemp, spattemp.density
```

Examples

```
data(burk) # Burkitt's Uganda lymphoma data
burkcas <- burk$cases

hlam1 <- LSCV.spattemp(burkcas) #~9 secs
hlam2 <- LSCV.spattemp(burkcas,tlim=c(400,5800))
hlam3 <- LSCV.spattemp(burkcas,start=c(7,400))
rbind(hlam1,hlam2,hlam3)

hlam1 <- LIK.spattemp(burkcas) #~3 secs
hlam2 <- LIK.spattemp(burkcas,tlim=c(400,5800))
hlam3 <- LIK.spattemp(burkcas,start=c(7,400))
rbind(hlam1,hlam2,hlam3)</pre>
```

LSCV.risk

Jointly optimal bandwidth selection for the spatial relative risk function

Description

Methods to find a jointly optimal, common case-control isotropic bandwidth for use in estimation of the fixed or adaptive kernel-smoothed relative risk function.

Usage

```
LSCV.risk(
  f,
  g = NULL,
  hlim = NULL,
```

```
hseq = NULL,
type = c("fixed", "adaptive"),
method = c("kelsall-diggle", "hazelton", "davies"),
resolution = 64,
edge = TRUE,
hp = NULL,
pilot.symmetry = c("none", "f", "g", "pooled"),
auto.optim = TRUE,
seqres = 30,
parallelise = NA,
verbose = TRUE,
...
)
```

Arguments

f

Either a pre-calculated object of class bivden representing the 'case' (numerator) density estimate, or an object of class ppp giving the observed case data. Alternatively, if f is ppp object with dichotomous factor-valued marks, the function treats the first level as the case data, and the second as the control data, obviating the need to supply g.

g

As for f, for the 'control' (denominator) density; this object must be of the same class as f. Ignored if, as stated above, f contains both case and control observations.

hlim

An optional vector of length 2 giving the limits of the optimisation routine with respect to the bandwidth. If unspecified, the function attempts to choose this automatically.

hseq

An optional increasing sequence of bandwidth values at which to manually evaluate the optimisation criterion. Used only in the case (!auto.optim && is.null(hlim)).

type

A character string; "fixed" (default) performs classical leave-one-out cross-validation for a jointly optimal fixed bandwidth. Alternatively, "adaptive" utilises multiscale adaptive kernel estimation (Davies & Baddeley, 2018) to run the cross-validation in an effort to find a suitable jointly optimal, common global bandwidth for the adaptive relative risk function. See 'Details'.

method

A character string controlling the selector to use. There are three types, based on either the mean integrated squared error (MISE) (Kelsall and Diggle, 1995; default – method = "kelsall-diggle"); a weighted MISE (Hazelton, 2008 – method = "hazelton"); or an approximation to the asymptotic MISE (Davies, 2013 – method = "davies"). See 'Details'.

resolution

Spatial grid size; the optimisation will be based on a [resolution \times resolution] density estimate.

edge

Logical value indicating whether to edge-correct the density estimates used.

hp

A single numeric value or a vector of length 2 giving the pilot bandwidth(s) to be used for estimation of the pilot densities for adaptive risk surfaces. Ignored if type = "fixed".

pilot.symmetry A character string used to control the type of symmetry, if any, to use for the bandwidth factors when computing an adaptive relative risk surface. See 'Details'. Ignored if type = "fixed". auto.optim Logical value indicating whether to automate the numerical optimisation using optimise. If FALSE, the optimisation criterion is evaluated over hseq (if supplied), or over a sequence of values controlled by hlim and segres. segres Optional resolution of an increasing sequence of bandwidth values. Only used if (!auto.optim && is.null(hseq)). parallelise Numeric argument to invoke parallel processing, giving the number of CPU cores to use when !auto.optim. Experimental. Test your system first using parallel::detectCores() to identify the number of cores available to you. Logical value indicating whether to provide function progress commentary. verbose Additional arguments such as dimz and trim to be passed to the internal calls to multiscale.density.

Details

Given the established preference of using a common bandwidth for both case and control density estimates when constructing a relative risk surface, This function calculates a 'jointly optimal', common isotropic LSCV bandwidth for the (Gaussian) kernel-smoothed relative risk function (case-control density-ratio). It can be shown that choosing a bandwidth that is equal for both case and control density estimates is preferable to computing 'separately optimal' bandwidths (Kelsall and Diggle, 1995). The user can choose to either calculate a common smoothing parameter for a fixed-bandwidth relative risk surface (type = "fixed"; default), or a common global bandwidth for an adaptive risk surface (type = "adaptive"). See further comments below.

- method = "kelsall-diggle": the function computes the common bandwidth which minimises the approximate mean integrated squared error (MISE) of the log-transformed risk surface (Kelsall and Diggle, 1995).
- method = "hazelton": the function minimises a weighted-by-control MISE of the (raw) relative risk function (Hazelton, 2008).
- method = "davies": the optimal bandwidth is one that minimises a crude plug-in approximation to the *asymptotic* MISE (Davies, 2013). Only possible for type = "fixed".

For jointly optimal, common global bandwidth selection when type = "adaptive", the optimisation routine utilises multiscale.density. Like LSCV.density, the leave-one-out procedure does not affect the pilot density, for which additional control is offered via the hp and pilot.symmetry arguments. The user has the option of obtaining a so-called *symmetric* estimate (Davies et al. 2016) via pilot.symmetry. This amounts to choosing the same pilot density for both case and control densities. By choosing "none" (default), the result uses the case and control data separately for the fixed-bandwidth pilots, providing the original asymmetric density-ratio of Davies and Hazelton (2010). By selecting either of "f", "g", or "pooled", the pilot density is calculated based on the case, control, or pooled case/control data respectively (using hp[1] as the fixed bandwidth). Davies et al. (2016) noted some beneficial practical behaviour of the symmetric adaptive surface over the asymmetric. (The pilot bandwidth(s), if not supplied in hp, are calculated internally via default use of LSCV.density, using the requested symmetric-based data set, or separately with respect to the case and control datasets f and g if pilot.symmetry = "none".)

Value

A single numeric value of the estimated bandwidth (if auto.optim = TRUE). Otherwise, a list of two numeric vectors of equal length giving the bandwidth sequence (as hs) and corresponding CV function value (as CV).

Warning

The jointly optimal bandwidth selector can be computationally expensive for large data sets and fine evaluation grid resolutions. The user may need to experiment with adjusting hlim to find a suitable minimum.

Author(s)

T. M. Davies

References

Davies, T. M. (2013), Jointly optimal bandwidth selection for the planar kernel-smoothed density-ratio, *Spatial and Spatio-temporal Epidemiology*, **5**, 51-65.

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Davies, T.M., Jones, K. and Hazelton, M.L. (2016), Symmetric adaptive smoothing regimens for estimation of the spatial relative risk function, *Computational Statistics & Data Analysis*, **101**, 12-28

Hazelton, M. L. (2008), Letter to the editor: Kernel estimation of risk surfaces without the need for edge correction, *Statistics in Medicine*, **27**, 2269-2272.

Kelsall, J.E. and Diggle, P.J. (1995), Kernel estimation of relative risk, *Bernoulli*, 1, 3-16.

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

Wand, M.P. and Jones, C.M., 1995. Kernel Smoothing, Chapman & Hall, London.

See Also

bivariate.density

Examples

```
data(pbc)
pbccas <- split(pbc)$case
pbccon <- split(pbc)$control
# FIXED (for common h)</pre>
```

multiscale.density 31

```
LSCV.risk(pbccas,pbccon,method="hazelton")
hcv <- LSCV.risk(pbccas,pbccon,method="davies",auto.optim=FALSE)
plot(hcv[,1],log(hcv[,2]));abline(v=hcv[which.min(hcv[,2]),1],col=2,lty=2)

# ADAPTIVE (for common h0)

LSCV.risk(pbccas,pbccon,type="adaptive")

# change pilot bandwidths used
LSCV.risk(pbccas,pbccon,type="adaptive",hp=c(0S(pbccas)/2,0S(pbccon)/2))

# specify pooled-data symmetric relative risk estimator
LSCV.risk(pbccas,pbccon,type="adaptive",hp=OS(pbc),pilot.symmetry="pooled")

# as above, for Hazelton selector
LSCV.risk(pbccas,pbccon,type="adaptive",method="hazelton")
LSCV.risk(pbccas,pbccon,type="adaptive",method="hazelton",hp=c(OS(pbccas)/2,OS(pbccon)/2))
LSCV.risk(pbccas,pbccon,type="adaptive",method="hazelton",hp=c(OS(pbccas)/2,OS(pbccon)/2))
LSCV.risk(pbccas,pbccon,type="adaptive",method="hazelton",hp=OS(pbc),pilot.symmetry="pooled")
```

multiscale.density

Multi-scale adaptive kernel density/intensity estimation

Description

Computes adaptive kernel estimates of spatial density/intensity using a 3D FFT for multiple global bandwidth scales.

Usage

```
multiscale.density(
  pp,
  h0,
  hp = NULL,
  h0fac = c(0.25, 1.5),
  edge = c("uniform", "none"),
  resolution = 128,
  dimz = 64,
  gamma.scale = "geometric",
  trim = 5,
  intensity = FALSE,
  pilot.density = NULL,
  xy = NULL,
  taper = TRUE,
  verbose = TRUE
)
```

32 multiscale.density

Arguments

рр	An object of class ppp giving the observed 2D data set to be smoothed.
h0	Reference global bandwidth for adaptive smoothing; numeric value > 0 . Multiscale estimates will be computed by rescaling this value as per h0fac.
hp	Pilot bandwidth (scalar, numeric > 0) to be used for fixed bandwidth estimation of the pilot density. If NULL (default), it will take on the value of h0. Ignored when pilot.density is supplied as a pre-defined pixel image.
h0fac	A numeric vector of length 2 stipulating the span of the global bandwidths in the multiscale estimates. Interpreted as a multiplicative factor on hø. See 'Details'.
edge	Character string dictating edge correction. "uniform" (default) corrects based on evaluation grid coordinate. Setting edge="none" requests no edge correction.
resolution	Numeric value > 0 . Resolution of evaluation grid in the spatial domain; the densities/intensities will be returned on a [resolution \times resolution] grid.
dimz	Resolution of z- (rescaled bandwidth)-axis in the trivariate convolution. Higher values increase precision of the multiscale estimates at a computational cost. See 'Details'.
gamma.scale	Scalar, numeric value > 0; controls rescaling of the variable bandwidths. Defaults to the geometric mean of the bandwidth factors given the pilot density (as per Silverman, 1986). See the documentation for bivariate.density.
trim	Numeric value > 0 ; controls bandwidth truncation for adaptive estimation. See the documentation for bivariate.density.
intensity	Logical value indicating whether to return an intensity estimate (integrates to the sample size over the study region), or a density estimate (default, integrates to 1).
pilot.density	An optional pixel image (class im) giving the pilot density to be used for calculation of the variable bandwidths in adaptive estimation, or a ppp.object giving the data upon which to base a fixed-bandwidth pilot estimate using hp. See the documentation for bivariate.density.
ху	Optional alternative specification of the spatial evaluation grid; matches the argument of the same tag in as.mask. If supplied, resolution is ignored.
taper	Logical value indicating whether to taper off the trivariate kernel outside the range of h0*h0fac in the scale space; see Davies & Baddeley (2018). Keep at the default TRUE if you don't know what this means.
verbose	Logical value indicating whether to print function progress.

Details

Davies & Baddeley (2018) investigated computational aspects of Abramson's (1982) adaptive kernel smoother for spatial (2D) data. This function is the implementation of the 3D convolution via a fast-Fourier transform (FFT) which allows simultaneous calculation of an adaptive kernel estimate at multiple global bandwidth scales.

These 'multiple global bandwidth scales' are computed with respect to rescaling a reference value of the global bandwidth passed to the h0 argument. This rescaling is defined by the range provided

multiscale.density 33

to the argument h0fac. For example, by default, the function will compute the adaptive kernel estimate for a range of global bandwidths between 0.25*h0 and 1.5*h0. The exact numeric limits are subject to discretisation, and so the returned valid range of global bandwidths will differ slightly. The exact resulting range following function execution is returned as the h0range element of the result, see 'Value' below.

The distinct values of global bandwidth used (which define the aforementioned h@range) and hence the total number of pixel images returned depend on both the width of the span h@fac and the discretisation applied to the bandwidth axis through dimz. Increasing this z-resolution will provide more pixel images and hence greater numeric precision, but increases computational cost. The returned pixel images that represent the multiscale estimates are stored in a named list (see 'Value'), whose names reflect the corresponding distinct global bandwidth. See 'Examples' for the easy way to extract these distinct global bandwidths.

The user can request an interpolated density/intensity estimate for any global bandwidth value within horange by using the multiscale.slice function, which returns an object of class bivden.

Value

An object of class "msden". This is very similar to a bivden object, with lists of pixel images in the z, him, and q components (instead of standalone images).

Z	A list of the resulting density/intensity estimates; each member being a pixel image object of class im. They are placed in increasing order of the discretised values of h0.
h0	A copy of the reference value of h0 used.
h0range	A vector of length 2 giving the actual range of global bandwidth values available (inclusive).
hp	A copy of the value of hp used.
h	A numeric vector of length equal to the number of data points, giving the bandwidth used for the corresponding observation in pp with respect to the reference global bandwidth h0.
him	A list of pixel images (class im), corresponding to z, giving the 'hypothetical' Abramson bandwidth at each pixel coordinate conditional upon the observed data and the global bandwidth used.
q	Edge-correction weights; list of pixel images corresponding to z if edge = "uniform", and NULL if edge = "none".
gamma	The numeric value of gamma.scale used in scaling the bandwidths.
geometric	The geometric mean of the untrimmed variable bandwidth factors. This will be identical to gamma if gamma.scale = "geometric" as per default.
pp	A copy of the ppp.object initially passed to the pp argument, containing the data that were smoothed.

Author(s)

T.M. Davies and A. Baddeley

34 multiscale.slice

References

Abramson, I. (1982). On bandwidth variation in kernel estimates — a square root law, *Annals of Statistics*, **10**(4), 1217-1223.

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

See Also

```
bivariate.density, multiscale.slice
```

Examples

```
data(chorley) # Chorley-Ribble data (package 'spatstat')
ch.multi <- multiscale.density(chorley,h0=1)
plot(ch.multi)

ch.pilot <- bivariate.density(chorley,h0=0.75) # with pre-defined pilot density
ch.multi2 <- multiscale.density(chorley,h0=1,pilot.density=ch.pilot$z)
plot(ch.multi2)

data(pbc)
# widen h0 scale, increase z-axis resolution
pbc.multi <- multiscale.density(pbc,h0=2,hp=1,h0fac=c(0.25,2.5),dimz=128)
plot(pbc.multi)</pre>
```

multiscale.slice

Slicing a multi-scale density/intensity object

Description

Takes slices of a multi-scale density/intensity estimate at desired global bandwidths

Usage

```
multiscale.slice(msob, h0, checkargs = TRUE)
```

Arguments

h0

msob	An object of class msden giving the multi-scale estimate from which to take
	slices

Desired global bandwidth(s); the density/intensity estimate corresponding to

which will be returned. A numeric vector. All values must be in the available

range provided by msob\$h0range; see 'Details'.

checkargs Logical value indicating whether to check validity of msob and h0. Disable only

if you know this check will be unnecessary.

multiscale.slice 35

Details

Davies & Baddeley (2018) demonstrate that once a multi-scale density/intensity estimate has been computed, we may take slices parallel to the spatial domain of the trivariate convolution to return the estimate at any desired global bandwidth. This function is the implementation thereof based on a multi-scale estimate resulting from a call to multiscale.density.

The function returns an error if the requested slices at h0 are not all within the available range of pre-computed global bandwidth scalings as defined by the h0range component of msob.

Because the contents of the msob argument, an object of class msden, are returned based on a discretised set of global bandwidth scalings, the function internally computes the desired surface as a pixel-by-pixel linear interpolation using the two discretised global bandwidth rescalings that bound each requested h0. (Thus, numeric accuracy of the slices is improved with an increase to the dimz argument of the preceding call to multiscale.density at the cost of additional computing time.)

Value

If h0 is scalar, an object of class bivden with components corresponding to the requested slice at h0. If h0 is a vector, a list of objects of class bivden.

Author(s)

T.M. Davies

References

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

See Also

```
multiscale.density, bivariate.density
```

Examples

```
data(chorley) # Chorley-Ribble data (package 'spatstat')
ch.multi <- multiscale.density(chorley,h0=1,h0fac=c(0.5,2))
available.h0(ch.multi)
ch.slices <- multiscale.slice(ch.multi,h0=c(0.7,1.1,1.6))
par(mfcol=c(2,3)) # plot each density and edge-correction surface
for(i in 1:3) { plot(ch.slices[[i]]$z); plot(ch.slices[[i]]$q) }</pre>
```

NS NS

Normal scale (NS) bandwidth selector

Description

Provides the asymptotically optimal fixed bandwidths for spatial or spatiotemporal normal densities based on a simple expression.

Usage

```
NS(
   pp,
   nstar = c("npoints", "geometric"),
   scaler = c("silverman", "IQR", "sd", "var")
)

NS.spattemp(
   pp,
   tt = NULL,
   nstar = "npoints",
   scaler = c("silverman", "IQR", "sd", "var")
)
```

Arguments

рр	An object of class ppp giving the observed 2D data to be smoothed.
nstar	Optional. Controls the value to use in place of the number of observations n in the normal scale formula. Either a character string, "npoints" (default) or "geometric" (only possible for NS), or a positive numeric value. See 'Details'.
scaler	Optional. Controls the value for a scalar representation of the spatial (and temporal for NS. spattemp) scale of the data. Either a character string, "silverman" (default), "IQR", "sd", or "var"; or a positive numeric value. See 'Details'.
tt	A numeric vector of equal length to the number of points in pp, giving the time corresponding to each spatial observation. If unsupplied, the function attempts to use the values in the marks attribute of the ppp.object in pp.

Details

These functions calculate scalar smoothing bandwidths for kernel density estimates of spatial or spatiotemporal data: the optimal values would minimise the asymptotic mean integrated squared error assuming normally distributed data; see pp. 46-48 of Silverman (1986). The NS function returns a single bandwidth for isotropic smoothing of spatial (2D) data. The NS. spattemp function returns two values — one for the spatial margin and another for the temporal margin, based on independently applying the normal scale rule (in 2D and 1D) to the spatial and temporal margins of the supplied data.

NS

Effective sample size The formula requires a sample size, and this can be minimally tailored via nstar. By default, the function simply uses the number of observations in pp: nstar = "npoints". Alternatively, the user can specify their own value by simply supplying a single positive numeric value to nstar. For NS (not applicable to NS.spattemp), if pp is a ppp.object with factor-valued marks, then the user has the option of using nstar = "geometric", which sets the sample size used in the formula to the geometric mean of the counts of observations of each mark. This can be useful for e.g. relative risk calculations, see Davies and Hazelton (2010).

Spatial (and temporal) scale The scaler argument is used to specify spatial (as well as temporal, in use of NS. spattemp) scale. For isotropic smoothing in the spatial margin, one may use the 'robust' estimate of standard deviation found by a weighted mean of the interquartile ranges of the x- and y-coordinates of the data respectively (scaler = "IQR"). Two other options are the raw mean of the coordinate-wise standard deviations (scaler = "sd"), or the square root of the mean of the two variances (scaler = "var"). A fourth option, scaler = "silverman" (default), sets the scaling constant to be the minimum of the "IQR" and "sd" options; see Silverman (1986), p. 47. In use of NS. spattemp the univariate version of the elected scale statistic is applied to the recorded times of the data for the temporal bandwidth. Alternatively, like nstar, the user can specify their own value by simply supplying a single positive numeric value to scaler for NS, or a numeric vector of length 2 (in the order of [<spatial scale>, <temporal scale>]) for NS. spattemp.

Value

A single numeric value of the estimated spatial bandwidth for NS, or a named numeric vector of length 2 giving the spatial bandwidth (as h) and the temporal bandwidth (as lambda) for NS. spattemp.

Warning

The NS bandwidth is an approximation, and assumes *that the target density is normal*. This is considered rare in most real-world applications. Nevertheless, it remains a quick and easy 'rule-of-thumb' method with which one may obtain a smoothing parameter. Note that a similar expression for the adaptive kernel estimator is not possible (Davies et al., 2018).

Author(s)

T.M. Davies

References

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Davies, T.M., Flynn, C.R. and Hazelton, M.L. (2018), On the utility of asymptotic bandwidth selectors for spatially adaptive kernel density estimation, *Statistics & Probability Letters* [in press].

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

Wand, M.P. and Jones, C.M., 1995. Kernel Smoothing, Chapman & Hall, London.

38 OS

Examples

```
data(pbc)
NS(pbc)
NS(pbc,nstar="geometric") # uses case-control marks to replace sample size
NS(pbc,scaler="var") # set different scalar measure of spread

data(burk)
NS.spattemp(burk$cases)
NS.spattemp(burk$cases,scaler="sd")
```

0S

Oversmoothing (OS) bandwidth selector

Description

Provides fixed bandwidths for spatial or spatiotemporal data based on the maximal smoothing (oversmoothing) principle of Terrell (1990).

Usage

```
OS(
   pp,
   nstar = c("npoints", "geometric"),
   scaler = c("silverman", "IQR", "sd", "var")
)

OS.spattemp(
   pp,
   tt = NULL,
   nstar = "npoints",
   scaler = c("silverman", "IQR", "sd", "var")
)
```

Arguments

pp	An object of class ppp giving the observed 2D data to be smoothed.
nstar	Optional. Controls the value to use in place of the number of observations n in the oversmoothing formula. Either a character string, "npoints" (default) or "geometric" (only possible for OS), or a positive numeric value. See 'Details'.
scaler	Optional. Controls the value for a scalar representation of the spatial (and temporal for OS. spattemp) scale of the data. Either a character string, "silverman" (default), "IQR", "sd", or "var"; or positive numeric value(s). See 'Details'.
tt	A numeric vector of equal length to the number of points in pp, giving the time corresponding to each spatial observation. If unsupplied, the function attempts to use the values in the marks attribute of the ppp.object in pp.

Details

These functions calculate scalar smoothing bandwidths for kernel density estimates of spatial or spatiotemporal data: the "maximal amount of smoothing compatible with the estimated scale of the observed data". See Terrell (1990). The OS function returns a single bandwidth for isotropic smoothing of spatial (2D) data. The OS. spattemp function returns two values – one for the spatial margin and another for the temporal margin, based on independently applying Terrell's (1990) rule (in 2D and 1D) to the spatial and temporal margins of the supplied data.

Effective sample size The formula requires a sample size, and this can be minimally tailored via nstar. By default, the function simply uses the number of observations in pp: nstar = "npoints". Alternatively, the user can specify their own value by simply supplying a single positive numeric value to nstar. For OS (not applicable to OS.spattemp), if pp is a ppp.object with factor-valued marks, then the user has the option of using nstar = "geometric", which sets the sample size used in the formula to the geometric mean of the counts of observations of each mark. This can be useful for e.g. relative risk calculations, see Davies and Hazelton (2010).

Spatial (and temporal) scale The scaler argument is used to specify spatial (as well as temporal, in use of OS. spattemp) scale. For isotropic smoothing in the spatial margin, one may use the 'robust' estimate of standard deviation found by a weighted mean of the interquartile ranges of the x- and y-coordinates of the data respectively (scaler = "IQR"). Two other options are the raw mean of the coordinate-wise standard deviations (scaler = "sd"), or the square root of the mean of the two variances (scaler = "var"). A fourth option, scaler = "silverman" (default), sets the scaling constant to be the minimum of the "IQR" and "sd" options; see Silverman (1986), p. 47. In use of OS. spattemp the univariate version of the elected scale statistic is applied to the recorded times of the data for the temporal bandwidth. Alternatively, like nstar, the user can specify their own value by simply supplying a single positive numeric value to scaler for OS, or a numeric vector of length 2 (in the order of [<spatial scale>, <temporal scale>]) for OS. spattemp.

Value

A single numeric value of the estimated spatial bandwidth for OS, or a named numeric vector of length 2 giving the spatial bandwidth (as h) and the temporal bandwidth (as lambda) for OS. spattemp.

Author(s)

T.M. Davies

References

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Terrell, G.R. (1990), The maximal smoothing principle in density estimation, *Journal of the American Statistical Association*, **85**, 470-477.

Examples

data(pbc)

40 pbc

```
OS(pbc)
OS(pbc,nstar="geometric") # uses case-control marks to replace sample size
OS(pbc,scaler="var") # set different scalar measure of spread

data(burk)
OS.spattemp(burk$cases)
OS.spattemp(burk$cases,scaler="sd")
```

pbc

Primary biliary cirrhosis data

Description

Data of the locations of 761 cases of primary biliary cirrhosis in several adjacent health regions of north-eastern England, along with 3020 controls representing the at-risk population, collected between 1987 and 1994. These data were first presented and analysed by Prince et al. (2001); subsequent analysis of these data in the spirit of sparr was performed in Davies and Hazelton (2010). Also included is the polygonal study region.

Format

pbc is a dichotomously marked ppp.object, with locations expressed in UK Ordnance Survey Coordinates (km).

Acknowledgements

The authors thank Prof. Peter Diggle for providing access to these data.

Source

Prince et al. (2001), The geographical distribution of primary biliary cirrhosis in a well-defined cohort, *Hepatology*, **34**, 1083-1088.

References

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

```
data(pbc)
summary(pbc)
plot(pbc)
```

plot.bivden 41

plot.bivden

Plotting sparr objects

Description

plot methods for classes bivden, stden, rrs, rrst and msden.

Usage

```
## S3 method for class 'bivden'
plot(
  Х,
 what = c("z", "edge", "bw"),
  add.pts = FALSE,
 auto.axes = TRUE,
 override.par = TRUE,
)
## S3 method for class 'msden'
plot(x, what = c("z", "edge", "bw"), sleep = 0.2, override.par = TRUE, ...)
## S3 method for class 'rrs'
plot(
  х,
  auto.axes = TRUE,
  tol.show = TRUE,
  tol.type = c("upper", "lower", "two.sided"),
  tol.args = list(levels = 0.05, lty = 1, drawlabels = TRUE),
)
## S3 method for class 'rrst'
plot(
 х,
  tselect = NULL,
  type = c("joint", "conditional"),
  fix.range = FALSE,
  tol.show = TRUE,
  tol.type = c("upper", "lower", "two.sided"),
  tol.args = list(levels = 0.05, lty = 1, drawlabels = TRUE),
  sleep = 0.2,
  override.par = TRUE,
  expscale = FALSE,
)
```

42 plot.bivden

```
## S3 method for class 'stden'
plot(
    x,
    tselect = NULL,
    type = c("joint", "conditional"),
    fix.range = FALSE,
    sleep = 0.2,
    override.par = TRUE,
    ...
)
```

Arguments

fix.range

X	An object of class bivden, stden, rrs, rrst, or msden.
what	A character string to select plotting of result ("z"; default); edge-correction surface ("edge"); or variable bandwidth surface ("bw").
add.pts	Logical value indicating whether to add the observations to the image plot using default points.
auto.axes	Logical value indicating whether to display the plot with automatically added x-y axes and an 'L' box in default styles.
override.par	Logical value indicating whether to override the existing graphics device parameters prior to plotting, resetting mfrow and mar. See 'Details' for when you might want to disable this.
• • •	Additional graphical parameters to be passed to plot.im, or in one instance, to plot.ppp (see 'Details').
sleep	Single positive numeric value giving the amount of time (in seconds) to Sys.sleep before drawing the next image in the animation.
tol.show	Logical value indicating whether to show pre-computed tolerance contours on the plot(s). The object x must already have the relevant <i>p</i> -value surface(s) stored in order for this argument to have any effect.
tol.type	A character string used to control the type of tolerance contour displayed; a test for elevated risk ("upper"), decreased risk ("lower"), or a two-tailed test (two.sided).
tol.args	A named list of valid arguments to be passed directly to contour to control the appearance of plotted contours. Commonly used items are levels, lty, lwd and drawlabels.
tselect	Either a single numeric value giving the time at which to return the plot, or a vector of length 2 giving an interval of times over which to plot. This argument must respect the stored temporal bound in x\$tlim, else an error will be thrown. By default, the full set of images (i.e. over the entire available time span) is plotted.
type	A character string to select plotting of joint/unconditional spatiotemporal estimate (default) or conditional spatial density given time.

Logical value indicating whether use the same color scale limits for each plot in the sequence. Ignored if the user supplies a pre-defined colourmap to the

plot.bivden 43

col argument, which is matched to ... above and passed to plot.im. See 'Examples'.

expscale

Logical value indicating whether to force a raw-risk scale. Useful for users wishing to plot a log-relative risk surface, but to have the raw-risk displayed on the colour ribbon.

Details

In all instances, visualisation is deferred to plot.im, for which there are a variety of customisations available the user can access via The one exception is when plotting observation-specific "diggle" edge correction factors—in this instance, a plot of the spatial observations is returned with size proportional to the influence of each correction weight.

When plotting a rrs object, a pre-computed *p*-value surface (see argument tolerate in risk) will automatically be superimposed at a significance level of 0.05. Greater flexibility in visualisation is gained by using tolerance in conjunction with contour.

An msden, stden, or rrst object is plotted as an animation, one pixel image after another, separated by sleep seconds. If instead you intend the individual images to be plotted in an array of images, you should first set up your plot device layout, and ensure override.par = FALSE so that the function does not reset these device parameters itself. In such an instance, one might also want to set sleep = 0.

Value

Plots to the relevant graphics device.

Author(s)

T.M. Davies

```
data(pbc)
data(fmd)
data(burk)

# 'bivden' object
pbcden <- bivariate.density(split(pbc)$case,h0=3,hp=2,adapt=TRUE,davies.baddeley=0.05,verbose=FALSE)
plot(pbcden)
plot(pbcden,what="bw",main="PBC cases\n variable bandwidth surface",xlab="Easting",ylab="Northing")

# 'stden' object
burkden <- spattemp.density(burk$cases,tres=128) # observation times are stored in marks(burk$cases)
plot(burkden,fix.range=TRUE,sleep=0.1) # animation
plot(burkden,tselect=c(1000,3000),type="conditional") # spatial densities conditional on each time

# 'rrs' object
pbcrr <- risk(pbc,h0=4,hp=3,adapt=TRUE,tolerate=TRUE,davies.baddeley=0.025,edge="diggle")
plot(pbcrr) # default
plot(pbcrr,tol.args=list(levels=c(0.05,0.01),lty=2:1,col="seagreen4"),auto.axes=FALSE)</pre>
```

44 print.bivden

print.bivden

Printing sparr objects

Description

print methods for classes bivden, stden, rrs, rrst and msden.

Usage

```
## S3 method for class 'bivden'
print(x, ...)

## S3 method for class 'msden'
print(x, ...)

## S3 method for class 'rrs'
print(x, ...)

## S3 method for class 'rrst'
print(x, ...)

## S3 method for class 'stden'
print(x, ...)
```

Arguments

x An object of class bivden, stden, rrs, rrst, or msden.
... Ignored.

Author(s)

T.M. Davies

rimpoly 45

rimpoly	Random point generation inside polygon

Description

Generates a random point pattern of n iid points with any specified distribution based on a pixel image and a corresponding polygonal window.

Usage

```
rimpoly(n, z, w = NULL, correction = 1.1, maxpass = 50)
```

Arguments

n	Number of points to generate.
Z	A pixel image of class im defining the probability density of the points, possibly unnormalised.
W	A polygonal window of class owin. See 'Details'.
correction	An adjustment to the number of points generated at the initial pass of the internal loop in an effort to minimise the total number of passes required to reach n points. See 'Details'.
maxpass	The maximum number of passes allowed before the function exits. If this is reached before n points are found that fall within w, a warning is issued.

Details

This function is a deliberate variant of rpoint (Baddeley et. al, 2015), to be accessed when the user desires a randomly generated point pattern based on a pixel image, but wants the window of the point pattern to be a corresponding irregular polygon, as opposed to a binary image mask (which, when converted to a polygon directly, gives jagged edges based on the union of the pixels). When the user specifies their own polygonal window, a while loop is called and repeated as many times as necessary (up to maxpass times) to find n points inside w (when w = NULL, then the aforementioned union of the pixels of z is used, obtained via as.polygonal(Window(z))). The loop is necessary because the standard behaviour of rpoint can (and often does) yield points that sit in corners of pixels which lie outside the corresponding w.

The correction argument is used to determine how many points are generated initially, which will be ceiling(correction*n); to minimise the number of required passes over the loop this is by default set to give a number slightly higher than the requested n.

An error is thrown if Window(z) and w do not overlap.

Value

An object of class ppp containing the n generated points, defined with the polygonal owin, w.

Author(s)

T.M. Davies

References

Baddeley, A., Rubak, E. and Turner, R. (2015) *Spatial Point Patterns: Methodology and Applications with R*, Chapman and Hall/CRC Press, UK.

Examples

```
data(pbc)
Y <- bivariate.density(pbc,h0=2.5,res=25)
# Direct use of 'rpoint':
A <- rpoint(500,Y$z)
npoints(A)
# Using 'rimpoly' without supplying polygon:
B \leftarrow rimpoly(500, Y$z)
npoints(B)
# Using 'rimpoly' with the original pbc polygonal window:
C <- rimpoly(500,Y$z,Window(Y$pp))</pre>
npoints(C)
oldpar <- par(mfrow=c(1,3))</pre>
plot(A,main="rpoint")
plot(B,main="rimpoly (no polygon supplied)")
plot(C,main="rimpoly (original polygon supplied)")
par(oldpar)
```

risk

Spatial relative risk/density ratio

Description

Estimates a *relative risk* function based on the ratio of two 2D kernel density estimates.

Usage

```
risk(
    f,
    g = NULL,
    log = TRUE,
    h0 = NULL,
    hp = h0,
    adapt = FALSE,
    shrink = FALSE,
```

```
shrink.args = list(rescale = TRUE, type = c("lasso", "Bithell"), lambda = NA),
  tolerate = FALSE,
  doplot = FALSE,
  pilot.symmetry = c("none", "f", "g", "pooled"),
  epsilon = 0,
  verbose = TRUE,
)
```

Arguments

f Either a pre-calculated object of class bivden representing the 'case' (numerator) density estimate, or an object of class ppp giving the observed case data. Alternatively, if f is ppp object with dichotomous factor-valued marks, the function treats the first level as the case data, and the second as the control data,

obviating the need to supply g.

As for f, for the 'control' (denominator) density; this object must be of the g same class as f. Ignored if, as stated above, f contains both case and control

observations.

Logical value indicating whether to return the (natural) log-transformed relative risk function as recommended by Kelsall and Diggle (1995a). Defaults to TRUE,

with the alternative being the raw density ratio.

A single positive numeric value or a vector of length 2 giving the global bandwidth(s) to be used for case/control density estimates; defaulting to a common oversmoothing bandwidth computed via OS on the pooled data using nstar =

"geometric" if unsupplied. Ignored if f and g are already bivden objects.

A single numeric value or a vector of length 2 giving the pilot bandwidth(s) to be used for fixed-bandwidth estimation of the pilot densities for adaptive risk surfaces. Ignored if adapt = FALSE or if f and g are already bivden objects.

A logical value indicating whether to employ adaptive smoothing for internally

estimating the densities. Ignored if f and g are already bivden objects.

A logical value indicating whether to compute the shrinkage estimator of Hazelton (2023). This is only possible for adapt=FALSE.

A named list of optional arguments controlling the shrinkage estimator. Possible entries are rescale (a logical value indicating whether to integrate to one with respect to the control distribution over the window); type (a character string stipulating the shrinkage methodology to be used, either the default "lasso" or the alternative "Bithell"); and lambda (a non-negative numeric value determining the degree of shrinkage towards uniform relative risk—when set to its

default NA, it is selected via cross-validation).

A logical value indicating whether to internally calculate a corresponding asymp-

totic p-value surface (for tolerance contours) for the estimated relative risk func-

tion. See 'Details'.

Logical. If TRUE, an image plot of the estimated relative risk function is produced using various visual presets. If additionally tolerate was TRUE, asymp-

totic tolerance contours are automatically added to the plot at a significance level

log

h0

hp

adapt

shrink

shrink.args

tolerate

doplot

of 0.05 for elevated risk (for more flexible options for calculating and plotting tolerance contours, see tolerance and tol.contour).

pilot.symmetry A character string used to control the type of symmetry, if any, to use for the

bandwidth factors when computing an adaptive relative risk surface. See 'De-

tails'. Ignored if adapt = FALSE.

epsilon A single non-negative numeric value used for optional scaling to produce ad-

ditive constant to each density in the raw ratio (see 'Details'). A zero value

requests no additive constant (default).

verbose Logical value indicating whether to print function progress during execution.

.. Additional arguments passed to any internal calls of bivariate.density for estimation of the requisite densities. Ignored if f and g are already bivden

objects.

Details

The relative risk function is defined here as the ratio of the 'case' density to the 'control' (Bithell, 1990; 1991). Using kernel density estimation to model these densities (Diggle, 1985), we obtain a workable estimate thereof. This function defines the risk function r in the following fashion:

```
r = (fd + epsilon*max(gd))/(gd + epsilon*max(gd)),
```

where fd and gd denote the case and control density estimates respectively. Note the (optional) additive constants defined by epsilon times the maximum of each of the densities in the numerator and denominator respectively (see Bowman and Azzalini, 1997). A more recent shrinkage estimator developed by Hazelton (2023) is also implemented.

The log-risk function rho, given by rho = log[r], is argued to be preferable in practice as it imparts a sense of symmetry in the way the case and control densities are treated (Kelsall and Diggle, 1995a;b). The option of log-transforming the returned risk function is therefore selected by default.

When computing adaptive relative risk functions, the user has the option of obtaining a so-called *symmetric* estimate (Davies et al. 2016) via pilot.symmetry. This amounts to choosing the same pilot density for both case and control densities. By choosing "none" (default), the result uses the case and control data separately for the fixed-bandwidth pilots, providing the original asymmetric density-ratio of Davies and Hazelton (2010). By selecting either of "f", "g", or "pooled", the pilot density is calculated based on the case, control, or pooled case/control data respectively (using hp[1] as the fixed bandwidth). Davies et al. (2016) noted some beneficial practical behaviour of the symmetric adaptive surface over the asymmetric.

If the user selects tolerate = TRUE, the function internally computes asymptotic tolerance contours as per Hazelton and Davies (2009) and Davies and Hazelton (2010). When adapt = FALSE, the reference density estimate (argument ref.density in tolerance) is taken to be the estimated control density. The returned pixel image of p-values (see 'Value') is interpreted as an upper-tailed test i.e. smaller p-values represent greater evidence in favour of significantly increased risk. For greater control over calculation of tolerance contours, use tolerance.

Value

An object of class "rrs". This is a named list with the following components:

rr	A pixel image of the estimated risk surface.
f	An object of class bivden used as the numerator or 'case' density estimate.
g	An object of class bivden used as the denominator or 'control' density estimate.
Р	Only included if tolerate = TRUE. A pixel image of the <i>p</i> -value surface for tolerance contours: NULL otherwise.

Author(s)

T.M. Davies

References

Bithell, J.F. (1990), An application of density estimation to geographical epidemiology, *Statistics in Medicine*, **9**, 691-701.

Bithell, J.F. (1991), Estimation of relative risk functions, Statistics in Medicine, 10, 1745-1751.

Bowman, A.W. and Azzalini A. (1997), *Applied Smoothing Techniques for Data Analysis: The Kernel Approach with S-Plus Illustrations*, Oxford University Press Inc., New York.

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Davies, T.M., Jones, K. and Hazelton, M.L. (2016), Symmetric adaptive smoothing regimens for estimation of the spatial relative risk function, *Computational Statistics & Data Analysis*, **101**, 12-28.

Diggle, P.J. (1985), A kernel method for smoothing point process data, *Journal of the Royal Statistical Society Series C*, **34**(2), 138-147.

Hazelton, M.L. (2023), Shrinkage estimators of the spatial relative risk function, *Submitted for publication*.

Hazelton, M.L. and Davies, T.M. (2009), Inference based on kernel estimates of the relative risk function in geographical epidemiology, *Biometrical Journal*, **51**(1), 98-109.

Kelsall, J.E. and Diggle, P.J. (1995a), Kernel estimation of relative risk, *Bernoulli*, 1, 3-16.

Kelsall, J.E. and Diggle, P.J. (1995b), Non-parametric estimation of spatial variation in relative risk, *Statistics in Medicine*, **14**, 2335-2342.

```
data(pbc)
pbccas <- split(pbc)$case
pbccon <- split(pbc)$control
h0 <- OS(pbc,nstar="geometric")

# Fixed (with tolerance contours)
pbcrr1 <- risk(pbccas,pbccon,h0=h0,tolerate=TRUE)

# Fixed shrinkage
pbcrr2 <- risk(pbccas,pbccon,h0=h0,shrink=TRUE,shrink.args=list(lambda=4))

# Asymmetric adaptive</pre>
```

50 SLIK.adapt

```
pbcrr3 <- risk(pbccas,pbccon,h0=h0,adapt=TRUE,hp=c(OS(pbccas)/2,OS(pbccon)/2),</pre>
               tolerate=TRUE,davies.baddeley=0.05)
# Symmetric (pooled) adaptive
pbcrr4 <- risk(pbccas,pbccon,h0=h0,adapt=TRUE,tolerate=TRUE,hp=OS(pbc)/2,</pre>
               pilot.symmetry="pooled",davies.baddeley=0.05)
# Symmetric (case) adaptive; from two existing 'bivden' objects
f <- bivariate.density(pbccas,h0=h0,hp=2,adapt=TRUE,pilot.density=pbccas,</pre>
                        edge="diggle",davies.baddeley=0.05,verbose=FALSE)
g <- bivariate.density(pbccon,h0=h0,hp=2,adapt=TRUE,pilot.density=pbccas,</pre>
                        edge="diggle",davies.baddeley=0.05,verbose=FALSE)
pbcrr5 <- risk(f,g,tolerate=TRUE,verbose=FALSE)</pre>
oldpar <- par(mfrow=c(2,2))</pre>
plot(pbcrr1,override.par=FALSE,main="Fixed")
plot(pbcrr2,override.par=FALSE,main="Fixed shrinkage")
plot(pbcrr3,override.par=FALSE,main="Asymmetric adaptive")
plot(pbcrr4,override.par=FALSE,main="Symmetric (pooled) adaptive")
par(oldpar)
```

SLIK.adapt

Simultaneous global/pilot likelihood bandwidth selection

Description

Isotropic global and pilot bandwidth selection for adaptive density/intensity based on likelihood cross-validation.

Usage

```
SLIK.adapt(
   pp,
   hold = TRUE,
   start = rep(OS(pp), 2),
   hlim = NULL,
   edge = TRUE,
   zero.action = c(-1, 0),
   optim.control = list(),
   parallelise = NULL,
   verbose = TRUE,
   ...
)
```

Arguments

pp

An object of class ppp giving the observed 2D data to be smoothed.

SLIK.adapt 51

hold Logical value indicating whether to hold the global and pilot bandwidths equal throughout the optimisation; defaults to TRUE. See 'Details'. A positively-valued numeric vector of length 2 giving the starting values to be start used for the global/pilot optimisation routine when hold = FALSE. Defaults to the oversmoothing bandwidth (OS) for both values; ignored when hold = TRUE. hlim An optional vector of length 2 giving the limits of the optimisation routine with respect to the bandwidth when hold = TRUE. If unspecified, the function attempts to choose this automatically. Ignored when hold = FALSE. Logical value indicating whether to edge-correct the density estimates used. edge A numeric vector of length 2, each value being either -1, 0 (default), 1 or 2 zero.action controlling how the function should behave in response to numerical errors at very small bandwidths, when such a bandwidth results in one or more zero or negative density values during the leave-one-out computations. See 'Details'. optim.control An optional list to be passed to the control argument of optim for further control over the numeric optimisation when hold = FALSE. See the documentation for optim for further details. Numeric argument to invoke parallel processing in the brute force leave-one-out parallelise calculations, giving the number of CPU cores to use. Experimental. Test your system first using parallel::detectCores() to identify the number of cores available to you. If NA (default), no parallelisation performed and a single loop is used. verbose Logical value indicating whether to provide function progress commentary. Additional arguments controlling density estimation for the internal calculations. Relevant arguments are resolution, gamma. scale, and trim. If unsupplied these default to 64, "geometric", and 5 respectively; see bivariate. density for a further explanation of these arguments.

Details

This function is a generalisation of LIK.density, and is used in attempts to simultaneously choose an optimal global and pilot bandwidth for adaptive kernel density estimates. Where LIK.density for adaptive estimates assumes the pilot density is held constant (and is not subject to the leave-one-out operations), this function allows the pilot bandwidth to vary alongside the global.

Thus, in contrast to LIK.density the internal leave-one-out operations now also affect the pilot estimation stage. Hence, the set of variable bandwidths changes as each point is left out. In turn, this means the leave-one-out operations must be computed by brute force, and this is computationally expensive.

Identifiability problems can sometimes arise when the global and pilot bandwidths are allowed to 'float freely' in the bivariate optimisation routine, which is the default behaviour of the function (with hold = FALSE). This can be curbed by setting hold = TRUE, which forces both the global and pilot to be held at the same value during optimisation. Doing this also has the beneficial side effect of turning the problem into one of univariate optimisation, thereby reducing total computational cost. Current work (Davies & Lawson, 2018) provides some empirical evidence that this strategy performs quite well in practice.

Like LSCV. density and LIK. density, the argument zero. action can be used to control the level of severity in response to small bandwidths that result (due to numerical error) in at least one density

52 SLIK.adapt

value being zero or less. When this argument is passed a vector of length 2, the first entry corresponds to the global bandwidth (and hence refers to checks of the final adaptive density estimate and its leave-one-out values) and the second to the pilot bandwidth (and hence checks the fixed-bandwidth pilot density and its leave-one-out values). Alternatively a single value may be supplied, which will be taken to be the same for both global and pilot. See the help page for LIK.density for an explanation of the four allowable values (-1, 0, 1, 2) for each component of this argument.

Value

A numeric vector of length 2 giving the likelihood-maximised global and pilot bandwidths.

Note

While theoretically valid, this is a largely experimental function. There is presently little in the literature to suggest how well this type of simultaneous global/pilot bandwidth selection might perform in practice. Current research efforts (Davies & Lawson, 2018) seek in part to address these questions.

Author(s)

T. M. Davies

References

Davies, T.M. and Lawson, A.B. (2018), An evaluation of likelihood-based bandwidth selectors for spatial and spatiotemporal kernel estimates, *Submitted for publication*.

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

Wand, M.P. and Jones, C.M., 1995. Kernel Smoothing, Chapman & Hall, London.

See Also

Functions for bandwidth selection in package spatstat: bw.diggle; bw.ppl; bw.scott; bw.frac.

```
data(pbc)
pbccas <- split(pbc)$case

SLIK.adapt(pbccas)
SLIK.adapt(pbccas,hold=TRUE)</pre>
```

spattemp.density 53

spattemp.density	Spatiotemporal kernel density estimation	

Description

Provides a fixed-bandwidth kernel estimate of continuous spatiotemporal data.

Usage

```
spattemp.density(pp, h = NULL, tt = NULL, lambda = NULL,
tlim = NULL, sedge = c("uniform", "none"), tedge = sedge,
sres = 128, tres = NULL, verbose = TRUE)
```

Arguments

рр	An object of class ppp giving the spatial coordinates of the observations to be smoothed. Possibly marked with the time of each event; see argument tt.
h	Fixed bandwidth to smooth the spatial margin. A numeric value > 0 . If unsupplied, the oversmoothing bandwidth is used as per OS.
tt	A numeric vector of equal length to the number of points in pp, giving the time corresponding to each spatial observation. If unsupplied, the function attempts to use the values in the marks attribute of the ppp.object in pp.
lambda	Fixed bandwidth to smooth the temporal margin; a numeric value > 0. If unsupplied, the function internally computes the Sheather-Jones bandwith using bw.SJ (Sheather & Jones, 1991).
tlim	A numeric vector of length 2 giving the limits of the temporal domain over which to smooth. If supplied, all times in tt must fall within this interval (equality with limits allowed). If unsupplied, the function simply uses the range of the observed temporal values.
sedge	Character string dictating spatial edge correction. "uniform" (default) corrects based on evaluation grid coordinate. Setting sedge="none" requests no edge correction.
tedge	As sedge, for temporal edge correction.
sres	Numeric value > 0. Resolution of the [sres \times sres] evaluation grid in the spatial margin.
tres	Numeric value > 0 . Resolution of the evaluation points in the temporal margin as defined by the tlim interval. If unsupplied, the density is evaluated at integer values between tlim[1] and tlim[2].
verbose	Logical value indicating whether to print a function progress bar to the console during evaluation.

54 spattemp.density

Details

This function produces a fixed-bandwidth kernel estimate of a single spatiotemporal density, with isotropic smoothing in the spatial margin, as per Fernando & Hazelton (2014). Estimates may be edge-corrected for an irregular spatial study window *and* for the bounds on the temporal margin as per tlim; this edge-correction is performed in precisely the same way as the "uniform" option in bivariate.density.

Specifically, for n trivariate points in space-time (pp, tt, tlim), we have

$$\hat{f}(x,t) = n^{-1} \sum_{i=1}^{n} h^{-2} \lambda^{-1} K((x-x_i)/h) L((t-t_i)/\lambda) / (q(x)q(t)),$$

where $x\in W\subset R^2$ and $t\in T\subset R$; K and L are the 2D and 1D Gaussian kernels controlled by fixed bandwidths h (h) and λ (lambda) respectively; and $q(x)=\int_W h^{-2}K((u-x)/h)du$ and $q(t)=\int_T \lambda^{-1}L((w-t)/\lambda)dw$ are optional edge-correction factors (sedge and tedge).

The above equation provides the *joint* or *unconditional* density at a given space-time location (x, t). In addition to this, the function also yields the *conditional* density at each grid time, defined as

$$\hat{f}(x|t) = \hat{f}(x,t)/\hat{f}(t),$$

where $\hat{f}(t) = n^{-1} \sum_{i=1}^n \lambda^{-1} L((t-t_i)/\lambda)/q(t)$ is the univariate kernel estimate of the temporal margin. Normalisation of the two versions $\hat{f}(x,t)$ and $\hat{f}(x|t)$ is the only way they differ. Where in the unconditional setting we have $\int_W \int_T \hat{f}(x,t) dt dx = 1$, in the conditional setting we have $\int_W \hat{f}(x|t) dx = 1$ for all t. See Fernando & Hazelton (2014) for further details and practical reasons as to why we might prefer one over the other in certain situations.

The objects returned by this function (see 'Value' below) are necessary for kernel estimation of spatiotemporal relative risk surfaces, which is performed by spattemp.risk.

Value

An object of class "stden". This is effectively a list with the following components:

Z	A named (by time-point) list of pixel images corresponding to the joint spatiotemporal density over space at each discretised time.
z.cond	A named (by time-point) list of pixel images corresponding to the conditional spatial density given each discretised time.
h	The scalar bandwidth used for spatial smoothing.
lambda	The scalar bandwidth used for temporal smoothing.
tlim	A numeric vector of length two giving the temporal bound of the density estimate.
spatial.z	A pixel image giving the overall spatial margin as a single 2D density estimate (i.e. ignoring time).
temporal.z	An object of class density giving the overall temporal margin as a single 1D density estimate (i.e. ignoring space).
qs	A pixel image giving the edge-correction surface for the spatial margin. NULL if sedge = "none".

spattemp.density 55

qt	A numeric vector giving the edge-correction weights for the temporal margin. NULL if tedge = "none".
pp	A ppp.object of the spatial data passed to the argument of the same name in the initial function call, with marks of the observation times.
tgrid	A numeric vector giving the discretised time grid at which the spatiotemporal density was evaluated (matches the names of z and z.cond).

Author(s)

T.M. Davies

References

Duong, T. (2007), ks: Kernel Density Estimation and Kernel Discriminant Analysis for Multivariate Data in R, *Journal of Statistical Software*, **21**(7), 1-16.

Fernando, W.T.P.S. and Hazelton, M.L. (2014), Generalizing the spatial relative risk function, *Spatial and Spatio-temporal Epidemiology*, **8**, 1-10.

Kelsall, J.E. and Diggle, P.J. (1995), Kernel estimation of relative risk, *Bernoulli*, 1, 3-16.

Sheather, S. J. and Jones, M. C. (1991), A reliable data-based bandwidth selection method for kernel density estimation. Journal of the Royal Statistical Society Series B, **53**, 683-690.

Silverman, B.W. (1986), *Density Estimation for Statistics and Data Analysis*, Chapman & Hall, New York.

See Also

```
bivariate.density, spattemp.risk, spattemp.slice
```

```
data(burk)
burkcas <- burk$cases

burkden1 <- spattemp.density(burkcas,tres=128)
summary(burkden1)

hlam <- LIK.spattemp(burkcas,tlim=c(400,5900),verbose=FALSE)
burkden2 <- spattemp.density(burkcas,h=hlam[1],lambda=hlam[2],tlim=c(400,5900),tres=256)
tims <- c(1000,2000,3500)
par(mfcol=c(2,3))
for(i in tims){
   plot(burkden2,i,override.par=FALSE,fix.range=TRUE,main=paste("joint",i))
   plot(burkden2,i,"conditional",override.par=FALSE,main=paste("cond.",i))
}</pre>
```

56 spattemp.risk

Description

Produces a spatiotemporal relative risk surface based on the ratio of two kernel estimates of spatiotemporal densities.

Usage

```
spattemp.risk(f, g, log = TRUE, tolerate = FALSE, finiteness = TRUE, verbose = TRUE)
```

Arguments

f	An object of class stden representing the 'case' (numerator) density estimate.
g	Either an object of class stden, or an object of class bivden for the 'control' (denominator) density estimate. This object must match the spatial (and temporal, if stden) domain of f completely; see 'Details'.
log	Logical value indicating whether to return the log relative risk (default) or the raw ratio.
tolerate	Logical value indicating whether to compute and return asymptotic p -value surfaces for elevated risk; see 'Details'.
finiteness	Logical value indicating whether to internally correct infinite risk (on the log-scale) to the nearest finite value to avoid numerical problems. A small extra computational cost is required.
verbose	Logical value indicating whether to print function progress during execution.

Details

Fernando & Hazelton (2014) generalise the spatial relative risk function (e.g. Kelsall & Diggle, 1995) to the spatiotemporal domain. This is the implementation of their work, yielding the generalised log-relative risk function for $x \in W \subset R^2$ and $t \in T \subset R$. It produces

$$\hat{\rho}(x,t) = \log(\hat{f}(x,t)) - \log(\hat{g}(x,t)),$$

where $\hat{f}(x,t)$ is a fixed-bandwidth kernel estimate of the spatiotemporal density of the cases (argument f) and $\hat{q}(x,t)$ is the same for the controls (argument g).

- When argument g is an object of class stden arising from a call to spattemp.density, the resolution, spatial domain, and temporal domain of this spatiotemporal estimate must match that of f exactly, else an error will be thrown.
- When argument g is an object of class bivden arising from a call to bivariate. density, it is assumed the 'at-risk' control density is static over time. In this instance, the above equation for the relative risk becomes $\hat{\rho} = \log(\hat{f}(x,t)) + \log|T| \log(g(x))$. The spatial density estimate in g must match the spatial domain of f exactly, else an error will be thrown.

spattemp.risk 57

• The estimate $\hat{\rho}(x,t)$ represents the joint or unconditional spatiotemporal relative risk over $W \times T$. This means that the raw relative risk $\hat{r}(x,t) = \exp \hat{\rho}(x,t)$ integrates to 1 with respect to the control density over space and time: $\int_W \int_T r(x,t)g(x,t)dtdx = 1$. This function also computes the **conditional** spatiotemporal relative risk at each time point, namely

$$\hat{\rho}(x|t) = \log \hat{f}(x|t) - \log \hat{g}(x|t),$$

where $\hat{f}(x|t)$ and $\hat{g}(x|t)$ are the conditional densities over space of the cases and controls given a specific time point t (see the documentation for spattemp.density). In terms of normalisation, we therefore have $\int_W r(x|t)g(x|t)dx=1$. In the case where \hat{g} is static over time, one may simply replace $\hat{g}(x|t)$ with $\hat{g}(x)$ in the above.

• Based on the asymptotic properties of the estimator, Fernando & Hazelton (2014) also define the calculation of tolerance contours for detecting statistically significant fluctuations in such spatiotemporal log-relative risk surfaces. This function can produce the required *p*-value surfaces by setting tolerate = TRUE; and if so, results are returned for both the unconditional (x,t) and conditional (xlt) surfaces. See the examples in the documentation for plot.rrst for details on how one may superimpose contours at specific *p*-values for given evaluation times *t* on a plot of relative risk on the spatial margin.

Value

An object of class "rrst". This is effectively a list with the following members:

rr	A named (by time-point) list of pixel images corresponding to the joint spatiotemporal relative risk over space at each discretised time.
rr.cond	A named list of pixel images corresponding to the conditional spatial relative risk given each discretised time.
P	A named list of pixel $images$ of the p -value surfaces testing for elevated risk for the joint estimate. If tolerate = FALSE, this will be NULL.
P.cond	As above, for the conditional relative risk surfaces.
f	A copy of the object f used in the initial call.
g	As above, for g.
tlim	A numeric vector of length two giving the temporal bound of the density estimate.

Author(s)

T.M. Davies

References

Fernando, W.T.P.S. and Hazelton, M.L. (2014), Generalizing the spatial relative risk function, *Spatial and Spatio-temporal Epidemiology*, **8**, 1-10.

See Also

spattemp.density, spattemp.slice, bivariate.density

58 spattemp.slice

Examples

```
data(fmd)
fmdcas <- fmd$cases</pre>
fmdcon <- fmd$controls</pre>
f <- spattemp.density(fmdcas,h=6,lambda=8) # stden object as time-varying case density
g <- bivariate.density(fmdcon,h0=6) # bivden object as time-static control density
rho <- spattemp.risk(f,g,tolerate=TRUE)</pre>
print(rho)
oldpar <- par(mfrow=c(2,3))</pre>
plot(rho$f$spatial.z,main="Spatial margin (cases)") # spatial margin of cases
plot(rho$f$temporal.z,main="Temporal margin (cases)") # temporal margin of cases
plot(rho$g$z,main="Spatial margin (controls)") # spatial margin of controls
plot(rho, tselect = 50, type = "conditional", tol.args = list(levels = c(0.05, 0.0001), linear list(levels = c(0.05, 0.0001), list(levels = c(0.05, 0.000
                             lty=2:1,lwd=1:2),override.par=FALSE)
plot(rho,tselect=100,type="conditional",tol.args=list(levels=c(0.05,0.0001),
                            lty=2:1,lwd=1:2),override.par=FALSE)
plot(rho, tselect=200, type="conditional", tol.args=list(levels=c(0.05, 0.0001), type="conditional", type="conditional
                            lty=2:1,lwd=1:2),override.par=FALSE)
par(oldpar)
```

spattemp.slice

Slicing a spatiotemporal object

Description

Takes slices of the spatiotemporal kernel density or relative risk function estimate at desired times

Usage

```
spattemp.slice(stob, tt, checkargs = TRUE)
```

Arguments

stob	An object of class stden or rrst giving the spatiotemporal estimate from which to take slices.
tt	Desired time(s); the density/risk surface estimate corresponding to which will be returned. This value must be in the available range provided by stob\$tlim; see 'Details'.
checkargs	Logical value indicating whether to check validity of stob and tt. Disable only if you know this check will be unnecessary.

spattemp.slice 59

Details

Contents of the stob argument are returned based on a discretised set of times. This function internally computes the desired surfaces as pixel-by-pixel linear interpolations using the two discretised times that bound each requested tt.

The function returns an error if any of the requested slices at tt are not within the available range of times as given by the tlim component of stob.

Value

A list of lists of pixel images, each of which corresponds to the requested times in tt, and are named as such.

If stob is an object of class stden:

z Pixel images of the joint spatiotemporal density corresponding to tt.

z.cond Pixel images of the conditional spatiotemporal density given each time in tt.

If stob is an object of class rrst:

rr	Pixel images of the joint spatiotemporal relative risk corresponding to tt.
rr.cond	Pixel images of the conditional spatiotemporal relative risk given each time in

tt.

P Only present if tolerate = TRUE in the preceding call to spattemp.risk. Pixel

images of the p-value surfaces for the joint spatiotemporal relative risk.

P. cond Only present if tolerate = TRUE in the preceding call to spattemp.risk. Pixel

images of the p-value surfaces for the conditional spatiotemporal relative risk.

Author(s)

T.M. Davies

References

Fernando, W.T.P.S. and Hazelton, M.L. (2014), Generalizing the spatial relative risk function, *Spatial and Spatio-temporal Epidemiology*, **8**, 1-10.

See Also

```
spattemp.density, spattemp.risk, bivariate.density
```

```
data(fmd)
fmdcas <- fmd$cases
fmdcon <- fmd$controls

f <- spattemp.density(fmdcas,h=6,lambda=8)
g <- bivariate.density(fmdcon,h0=6)
rho <- spattemp.risk(f,g,tolerate=TRUE)</pre>
```

60 summary.bivden

```
f$tlim # requested slices must be in this range
# slicing 'stden' object
f.slice1 <- spattemp.slice(f,tt=50) # evaluation timestamp</pre>
f.slice2 <- spattemp.slice(f,tt=150.5) # interpolated timestamp</pre>
oldpar <- par(mfrow=c(2,2))</pre>
plot(f.slice1$z$'50')
plot(f.slice1$z.cond$'50')
plot(f.slice2$z$'150.5')
plot(f.slice2$z.cond$'150.5')
par(oldpar)
# slicing 'rrst' object
rho.slices <- spattemp.slice(rho,tt=c(50,150.5))</pre>
oldpar <- par(mfrow=c(2,2))</pre>
plot(rho.slices$rr$'50');tol.contour(rho.slices$P$'50',levels=0.05,add=TRUE)
plot(rho.slices$rr$'150.5'); tol.contour(rho.slices$P$'150.5',levels=0.05,add=TRUE)
plot(rho.slices$rr.cond$'50');tol.contour(rho.slices$P.cond$'50',levels=0.05,add=TRUE)
plot(rho.slices$rr.cond$'150.5');tol.contour(rho.slices$P.cond$'150.5',levels=0.05,add=TRUE)
par(oldpar)
```

summary.bivden

Summarising sparr objects

Description

summary methods for classes bivden, stden, rrs, rrst and msden.

Usage

```
## S3 method for class 'bivden'
summary(object, ...)

## S3 method for class 'msden'
summary(object, ...)

## S3 method for class 'rrs'
summary(object, ...)

## S3 method for class 'rrst'
summary(object, ...)

## S3 method for class 'stden'
summary(object, ...)
```

tol.classify 61

Arguments

object An object of class bivden, stden, rrs, rrst, or msden.

... Ignored.

Author(s)

T.M. Davies

tol.classify Classification by p-value surfaces

Description

Classifies observed case/control points according to an estimated *p*-value surface.

Usage

```
tol.classify(rs, cutoff = 0.05, pim = NULL, ...)
```

Arguments

rs	An object of class rrs giving the estimated relative risk function of the case-control points to be classified.
cutoff	A numeric value between 0 and 1, defining the cutoff p-value used to classify points; defaults to 0.05.
pim	A pixel image defining the p-value surface with respect to which the observations are to be classified. Typically the result of a call to tolerance. Ignored if rs possesses a non-NULL P component (which takes precedence).
•••	Arguments to be passed to tolerance in order to compute the desired p-value surface in the event neither rs\$P nor pim exist. Ignored otherwise.

Details

This function takes in a relative risk surface computed with risk and corresponding p-value surface (the latter used for drawing tolerance contours), and attempts to classify both the case and control points as either falling within or without contours drawn at a level of cutoff. Points that fall 'inside' the contours are deemed to be associated with p-values less than or equal to cutoff and hence are usually interpreted as being in spatial areas of significant risk. This is useful for identifying characteristics of points that fall inside 'pockets of significance' as delineated by tolerance contours.

Upon execution, the function first inspects the rs object to determine whether it possesses a P component (i.e. an internally computed p-value surface provided when risk is called with optional argument tolerate=TRUE). If it exists, this is used. If not, the function then looks to see if the pim argument has been supplied. If it has, it must be a pixel image compatible with the risk surface in rs\$rr. If neither rs\$P nor pim is present, the function internally calls tolerance with arguments supplied to . . . to produce the desired surface.

62 tol.classify

The return object is a list that splits each of the case and control ppp data objects (these are stored as rs\$f\$pp and rs\$g\$pp) in the originally supplied risk surface object) into two constituent ppp objects – one comprising the points inside the cutoff contours (fin and gin), the other for those points outside the cutoff contours (fout and gout). In addition, the index values of the original data objects rs\$f\$pp and rs\$g\$pp that correspond to the points in fin and gin are provided as numeric vectors (findex and gindex). These objects are useful if you need to cross-reference data-specific characteristics from some other (corresponding) data set.

Further supplied in the returned list are quantities describing the overall classification structure (pcmask), as well as contour-specific identification and classification (finsplit, ginsplit, pcpolys). The pcpolys object can be plotted to illustrate the unique contour IDs with tol.classplot.

Value

A list of ten components:

fin	Point pattern of 'case' observations classified as being inside the cutoff contours of the p-value surface. An object of class ppp.
fout	Point pattern of 'case' observations falling outside the cutoff contours of the p-value surface. An object of class ppp.
gin	As fin, for the control points.
gout	As fout, for the control points.
findex	Numeric vector giving the raw index values of the original pattern of cases which provide fin.
gindex	As findex, for the controls.
finsplit	A list of the indexes in findex, with separate members splitting up the indexes of case observations as falling inside each unique tolerance contour.
ginsplit	As ginsplit, for the controls.
pcmask	The classification object of class owin. This is a pixel image mask derived from pim and cutoff.
pcpolys	A list of the same length as finsplit and ginsplit, identifying each unique contour as a polygonal owin. The order of these objects in the list correspond to the membership of finsplit and ginsplit. Use tol.classplot on this component to plot the classification indexes.

Author(s)

T. M. Davies

References

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Hazelton, M.L. and Davies, T.M. (2009), Inference based on kernel estimates of the relative risk function in geographical epidemiology, *Biometrical Journal*, **51**(1), 98-109.

Kelsall, J.E. and Diggle, P.J. (1995), Kernel estimation of relative risk, *Bernoulli*, 1, 3-16.

tol.classplot 63

Examples

```
data(pbc)
pbccas <- split(pbc)$case</pre>
pbccon <- split(pbc)$control</pre>
h0 <- OS(pbc,nstar="geometric")</pre>
pbcrr <- risk(pbccas,pbccon,h0=h0,tolerate=TRUE)</pre>
pbcclass <- tol.classify(pbcrr)</pre>
## Not run:
plot(pbcrr)
points(pbcclass$fin,col="red",pch=3,cex=0.5)
points(pbcclass$fout,col="seagreen4",cex=0.5)
chrr <- risk(chorley,h0=0.7,tolerate=TRUE)</pre>
chclass <- tol.classify(chrr,cutoff=0.4)</pre>
plot(chrr,tol.args=list(levels=0.4))
for(i in 1:length(chclass$finsplit)){
   points(chrr$f$pp[chclass$finsplit[[i]]],col=i,pch=19)
}
## End(Not run)
```

tol.classplot

Plot tolerance contour classification scheme

Description

Permits illustration of the uniquely identified tolerance contour regions arising from a call to tol.classify.

Usage

```
tol.classplot(pcpolys, add = FALSE, annotate = TRUE, ...)
```

Arguments

pcpolys	A list of polygonal windows, each of class owin. This will almost always be the pcpolys component of the object returned by a call to tol.classify.
add	A logical value indicating whether to add the unique regions to an existing plot (see 'Examples').
annotate	A logical value indicating whether to annotate each unique region with its identifying number (which will correspond to the uniquely split/classified points in a corresponding call to tol.classify).
• • •	Additional arguments to be passed to text to control the appearance of the annotations when annotate=TRUE.

64 tol.classplot

Details

The tol.classify function permits identification of individual significance regions (that is, the tolerance contours). In turn, tol.classplot may be used to visualise these regions optionally annotated by their unique identification number to better understand the region-specific classifications of the case and control points.

Value

Plots to the relevant graphics device.

Author(s)

T. M. Davies

References

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Hazelton, M.L. and Davies, T.M. (2009), Inference based on kernel estimates of the relative risk function in geographical epidemiology, *Biometrical Journal*, **51**(1), 98-109.

Kelsall, J.E. and Diggle, P.J. (1995), Kernel estimation of relative risk, Bernoulli, 1, 3-16.

```
## Not run:
chrr <- risk(chorley,h0=0.7,tolerate=TRUE)
chclass <- tol.classify(chrr,cutoff=0.4)

oldpar <- par(mfrow=c(1,3))
#
plot(chrr,tol.args=list(levels=0.4))
tol.classplot(chclass$pcpolys)
plot(Window(chorley))
axis(1)
axis(2)
box(bty="1")
tol.classplot(chclass$pcpolys,add=TRUE,col=2,font=2,cex=1.5)
#
par(oldpar)

## End(Not run)</pre>
```

tol.contour 65

ance contours

Description

Draw contours based on a *p*-value matrix.

Usage

```
tol.contour(pim, test = c("upper", "lower", "two-sided"), ...)
```

Arguments

pim	A pixel image of p -values, typically obtained from a call to tolerance, computed with respect to a test for elevated risk.
test	An optional character string giving the type of manipulation to be applied to the <i>p</i> -values, corresponding to a test for significantly elevated risk ("upper"; default); for reduced risk ("lower"); or for both ("two-sided").
	Additional arguments to be passed to contour. Commonly used options include add (to superimpose the contours upon an existing plot); levels (to control the specific significance levels at which to delineate the p -values); and lty or lwd for aesthetics.

Details

Note that no checks on the numeric content of pim are made. The function assumes the pixel image of *p*-values in pim is supplied with respect to an upper-tailed test for elevated risk (this is exactly the way the *p*-value surface is returned when tolerance is used). This is important if one makes subsequent use of test, which manipulates the *p*-values to draw at desired significance levels.

Value

Opens a new graphics device and displays a contour plot if add = FALSE, otherwise adds the contours to the plot in the existing active graphics device.

Author(s)

T. M. Davies

Examples

See ?tolerance

66 tolerance

tolerance

Tolerance by p-value surfaces

Description

Calculates a *p*-value surface based on asymptotic theory or Monte-Carlo (MC) permutations describing the extremity of risk given a fixed or adaptive kernel-smoothed density-ratio, allowing the drawing of *tolerance contours*.

Usage

```
tolerance(
   rs,
   method = c("ASY", "MC"),
   ref.density = NULL,
   beta = 0.025,
   ITER = 100,
   parallelise = NULL,
   verbose = TRUE,
   ...
)
```

Arguments

rs	An object of class rrs giving the estimated relative risk function for which to
	aslaulata tha m value aumfaas

calculate the *p*-value surface.

method A character string specifying the method of calculation. "ASY" (default) in-

structs the function to compute the p-values using asymptotic theory. "MC" com-

putes the values by random permutations of the data. See 'Details'.

ref.density Required if rs is based on fixed-bandwidth estimates of the case and control den-

sities and method = "ASY". Either a pixel image or an object of class bivden giving the reference density to use in asymptotic formulae. May be unnormalised.

Ignored if rs is based on adaptive kernel estimates or if method = "MC".

beta A numeric value $0 < \mathsf{beta} < 1$ giving the fineness of the adaptive bandwidth

partitioning to use for calculation of the required quantities for asymptotic adaptive p-value surfaces. Smaller values provide more accurate bandwidth bins at the cost of additional computing time, see Davies and Baddeley (2018); the default is sensible in most cases. Ignored if rs is based on fixed-bandwidth kernel

estimates.

ITER Number of iterations for the Monte-Carlo permutations. Ignored if method =

"ASY".

parallelise Numeric argument to invoke parallel processing, giving the number of CPU

cores to use when method = "MC". Experimental. Test your system first using parallel::detectCores() to identify the number of cores available to you.

verbose Logical value indicating whether to print function progress during execution.

tolerance 67

Additional arguments to be passed to risk when method = "MC". While most information needed for the MC repetitions is implicitly gleaned from the object passed to rs, this ellipsis is typically used to set the appropriate epsilon and pilot.symmetry values for the internal calls to risk.

Details

. . .

This function implements developments in Hazelton and Davies (2009) (fixed) and Davies and Hazelton (2010) (adaptive) to compute pointwise *p*-value surfaces based on asymptotic theory of kernel-smoothed relative risk surfaces. Alternatively, the user may elect to calculate the *p*-value surfaces using Monte-Carlo methods (see Kelsall and Diggle, 1995). Superimposition upon a plot of the risk surface contours of these *p*-values at given significance levels (i.e. "tolerance contours") can be an informative way of exploring the statistical significance of the extremity of risk across the defined study region.

Implementation of the Monte-Carlo method simply involves random allocation of case/control marks and re-estimation of the risk surface ITER times, against which the original estimate is compared. While not dependent on asymptotic theory, it is computationally expensive, and it has been suggested that it might have some undesirable practical consequences in certain settings (Hazelton and Davies, 2009). When performing the MC simulations, the same global (and pilot, if necessary) bandwidths and edge-correction regimens are employed as were used in the initial density estimates of the observed data. With regard to arguments to be passed to internal calls of risk, the user should take care to use . . . to set the epsilon value to match that which was used in creation of the object passed to rs (if this was set to a non-default value). Furthermore, if performing MC simulations for the adaptive relative risk function, the function borrows the value of the beta argument to speed things up via partitioning, and the user should additionally access . . . to set the same pilot.symmetry value as was used for creation of the object passed to rs, in the same way as for any non-default use of epsilon. This will ensure the simulations are all performed under the same conditions as were used to estimate the original risk function.

Value

A pixel image of the estimated p-value surface.

Note

The returned *p*-values are geared so that "smallness" corresponds to statistical significance of elevated risk, that is, an upper-tailed test. The complement of the *p*-values will yield significance of reduced risk; a lower-tailed test. When using tol.contour, the user can control what type of contours to display.

Author(s)

T. M. Davies

References

Davies, T.M. and Baddeley A. (2018), Fast computation of spatially adaptive kernel estimates, *Statistics and Computing*, **28**(4), 937-956.

68 tolerance

Davies, T.M. and Hazelton, M.L. (2010), Adaptive kernel estimation of spatial relative risk, *Statistics in Medicine*, **29**(23) 2423-2437.

Davies, T.M., Jones, K. and Hazelton, M.L. (2016), Symmetric adaptive smoothing regimens for estimation of the spatial relative risk function, *Computational Statistics & Data Analysis*, **101**, 12-28.

Hazelton, M.L. and Davies, T.M. (2009), Inference based on kernel estimates of the relative risk function in geographical epidemiology, *Biometrical Journal*, **51**(1), 98-109.

Kelsall, J.E. and Diggle, P.J. (1995), Kernel estimation of relative risk, Bernoulli, 1, 3-16.

```
data(pbc)
h0 <- LSCV.risk(pbc,method="hazelton");h0
pbccas <- split(pbc)[[1]]</pre>
pbccon <- split(pbc)[[2]]</pre>
# ASY
riskfix <- risk(pbc,h0=h0)
fixtol1 <- tolerance(riskfix,ref.density=density(pbc,OS(pbc)))</pre>
riskada <- risk(pbc,h0=h0,adapt=TRUE,hp=NS(pbc),pilot.symmetry="pooled",davies.baddeley=0.025)
adatol1 <- tolerance(riskada)</pre>
oldpar <- par(mfrow=c(1,2))</pre>
plot(riskfix)
tol.contour(fixtol1,levels=c(0.1,0.05,0.01),lty=3:1,add=TRUE)
plot(riskada)
tol.contour(adatol1,levels=c(0.1,0.05,0.01),lty=3:1,add=TRUE)
par(oldpar)
# MC
fixtol2 <- tolerance(riskfix,method="MC",ITER=200)</pre>
adatol2 <- tolerance(riskada, method="MC",ITER=200, parallelise=2) # ~90secs with parallelisation
oldpar <- par(mfrow=c(1,2))</pre>
plot(riskfix)
tol.contour(fixtol2,levels=c(0.1,0.05,0.01),lty=3:1,add=TRUE)
plot(riskada)
tol.contour(adatol2,levels=c(0.1,0.05,0.01),lty=3:1,add=TRUE)
par(oldpar)
```

Index

* data	LIK.density, 4, 22, 51, 52
burk, 19	LIK.spattemp, 5, 25
fmd, 21	LSCV.density, 4, 29, 51
pbc, 40	LSCV.density(LIK.density), 22
* package	LSCV.risk, 4, 27
sparr-package, 2	LSCV.spattemp, 5, 18
	LSCV.spattemp (LIK.spattemp), 25
as.mask, 10, 32	,
available.h0,8	marks, 17, 21, 26, 28, 36–39, 47, 53, 55
	msden, 8, 34, 35, 41–44, 60, 61
bivariate.density, 4, 9, 13, 15, 16, 23, 30,	msden (multiscale.density), 31
32, 34, 35, 48, 51, 54–57, 59	multiscale.density, 4, 8, 14-16, 23, 24, 29,
bivden, 14, 15, 28, 33, 35, 41, 42, 44, 47–49,	31, 35
56, 60, 61, 66	multiscale.slice, 4, 8, 33, 34, 34
bivden (bivariate.density), 9	martiseare. 311ee, 7, 0, 33, 37, 31
B00T.density, 4, 13	NS, 4, 36
BOOT. spattemp, <i>5</i> , 16, 27	NS.spattemp, 5
burk, 3, 19	no. Spacecinp, 5
burkitt, 19	optim, 17, 18, 26, 51
bw.diggle, 25, 52	optimise, 23, 29
	0S, 4, 13, 14, 16, 17, 38, 47, 51, 53
bw.frac, 25, 52	OS. spattemp, 5
bw.ppl, 25, 52	owin, 3, 19, 45, 62, 63
bw.scott, 25, 52	OWIN, 3, 17, 43, 02, 03
bw.SJ, <i>53</i>	parallel, 5
-h1 2	pbc, 3, 40
chorley, 3	plot.bivden, <i>4</i> , 41
colourmap, 42	plot.im, 42, 43
contour, <i>42</i> , <i>43</i> , <i>65</i>	plot.mm, 42, 43 plot.msden, 4
1 54	·
density, 54	plot.msden(plot.bivden), 41
density.ppp, 10, 11	plot.ppp, 42
doParallel, 5	plot.rrs, 4
00. 00	plot.rrs(plot.bivden), 41
fft, 20	plot.rrst, 5, 57
fft2d, 20	plot.rrst(plot.bivden), 41
fmd, 3, 21	plot.stden, 5
foreach, <i>5</i> , <i>11</i>	plot.stden(plot.bivden), 41
	points, 42
im, 3, 5, 10–12, 32, 33, 45, 48, 49, 54, 57, 59,	ppp, 5, 9, 13, 17, 19, 21, 23, 26, 28, 32, 36, 38,
61, 65–67	45, 47, 50, 53, 62

70 INDEX

```
ppp.object, 3, 10, 12, 17, 26, 32, 33, 36-40,
         53, 55
print.bivden, 4, 44
print.msden, 4
print.msden (print.bivden), 44
print.rrs, 4
print.rrs (print.bivden), 44
print.rrst, 5
print.rrst (print.bivden), 44
print.stden, 5
print.stden(print.bivden), 44
rimpoly, 3, 45
risk, 4, 43, 46, 61, 67
rpoint, 3, 45
rrs, 41–44, 60, 61, 66
rrs (risk), 46
rrst, 41–44, 58–61
rrst(spattemp.risk), 56
SLIK.adapt, 4, 24, 25, 50
sparr, 40
sparr (sparr-package), 2
sparr-package, 2
spatstat, 3, 5
spattemp.density, 5, 18, 27, 53, 56, 57, 59
spattemp.risk, 5, 54, 55, 56, 59
spattemp.slice, 5, 55, 57, 58
stden, 17, 18, 41-44, 56, 58-61
stden (spattemp.density), 53
summary.bivden, 4, 60
summary.msden, 4
summary.msden(summary.bivden), 60
summary.rrs, 4
summary.rrs (summary.bivden), 60
summary.rrst, 5
summary.rrst(summary.bivden), 60
summary.stden, 5
summary.stden(summary.bivden), 60
Sys.sleep, 42
text, 63
tol.classify, 4, 61, 63, 64
tol.classplot, 4, 62, 63
tol.contour, 4, 48, 65, 67
tolerance, 4, 43, 48, 61, 65, 66
```